

Substitution patterns and price response for plant-based meat alternatives

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Abstract

Efforts to promote sustainable resource use through reduced meat consumption face challenges as global meat consumption persists. The resistance may be attributed to the lower sales price of meat compared to most plant-based meat alternatives (PB-MAs). Addressing this, our research delves into the pivotal question of which PB-MAs resonate most with consumers and how pricing affects demand. In a hypothetical restaurant context, we conducted 2 representative studies among 2,126 individuals in the U.S. to scrutinize preferences for meat, analog, semi-analog, and non-analog burgers. First, in a survey, we assessed rankings of the four burgers, alongside evaluating participants' genuine consideration of these choices to discern a diverse preference distribution. Subsequently, in an experiment, we examined the influence of prices on participants' consideration and choice of PB-MAs, thereby capturing both phases of the decision-making process. Our survey shows that meat has considerably higher utility and consumer preference than all PB-MAs on average, but we also find substantial heterogeneity (i.e., some consumers prefer PB-MAs over meat). In the experiment, we establish that there is a negative association between the consideration of meat and PBMA burgers, though consideration of any one PBMA is positively associated with considering other PB-MAs. A noteworthy increase in consideration and choice is observed when prices of PB-MAs are reduced, while changing the price of the meat burger only has minimal effect on demand. Such findings underscore the importance of affordability beyond price parity in catalyzing the shift towards plant-based diets.

Keywords: Plant-based meat; food decision making; sustainability; price elasticity

Classification: Social Sciences, Economic Sciences

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Significance Statement

Diminishing meat intake in wealthier nations is crucial for environmental preservation and public health. Plant-based meat alternatives (PBMA) can help accomplish this goal, yet limited research explores complex substitution patterns and the influence of price adjustments. Further, we are aware of no studies examining the impact of price variation at the consideration and choice stage. Through robust survey and experimental research, we demonstrate that meat has considerably higher utility than PBMA on average. Still, demand for PBMA is sizeable, especially among certain consumer types. Encouragingly, individuals who begin to consider one PBMA type are more likely to weigh others. Lowering PBMA prices emerges as a strategic lever, offering a potential avenue to achieve the outcome of heightened consideration and choice of sustainable food.

Main Text

Introduction

Scientists agree that decreasing consumption of meat, particularly in wealthier nations, is an effective means towards sustainable use of global resources [1, 2]. Yet global meat consumption shows few signs of decline [3]. To some extent, this can be attributed to insufficient supply and/or promotion of attractive plant-based meat alternatives (hereafter PBMA). Accordingly, a central question is which kind of PBMA will resonate the most with consumers and how affordability can facilitate demand.

In a systematic review on consumer acceptance of alternative proteins, Onwezen and colleagues [4] concluded that consumers are more willing to accept plant-based novel proteins than animal-based novel proteins, such as insects and cultured meat. Circus and Robison [5] found that consumers with high (vs. low) meat attachment are more willing to eat meat alternatives but remain skeptical overall. A surprising limitation of previous research is the undifferentiated view of PBMA. The majority of studies lump together meat alternatives

or compare one PBMA with its meat counterpart [4, 6]. Importantly, plant-based protein has been around for thousands of years [7]. Early vegetarian proteins, such as tofu and tempeh, were mainly consumed by vegetarians [8]. These foods bore little resemblance to their meat counterparts, whereas recent advancements have focused on creating plant-based meat analogs that closely mimic the taste and texture of meat proteins, using innovative technologies and ingredients like pea protein and heme [9, 10]. Well-known examples of analogs are Beyond Meat and the Impossible Burger; such products are now commonplace in grocery stores as well as fast-food chains such as Burger King and Shake Shack. Recently, British consumer goods giant Unilever partnered with German Düzgün Group to develop a plant-based Döner kebab skewer for sale in kebab shops [11].

Despite any technological advancements, it is debated whether meat-mimicking PBMAAs are a fad or the future [12]. During the coronavirus pandemic in 2020, for example, Beyond Meat reported a series of better-than-expected quarterly results, but revenue has plummeted ever since. The up and down suggests general interest in meat analogs as novelty-seeking behavior [13] that could not be sustained, perhaps because not all consumers liked the product at the time. Favorable characteristics of meat analogs, such as similar sensory experience and familiar preparation [6], notwithstanding, a study among 2,497 Swedish adult consumers found that lightly processed legumes were seen as more attractive than legume-based meat analogs [14]. Despite major interest in the topic of PBMAAs, with more than 1,800 published articles in the Web of Science Core Collection as of July, 2024 [15], there is limited empirical evidence on how consumers perceive the various existing PBMAAs. In response, in this research, we consider three types of PBMAAs: (i) analog, (ii) semi-analog, and (iii) anon-analog. Analog PBMAAs try to mimic meat in every possible way, whereas semi-analog PBMAAs can be seen as first-generation alternatives that, like a veggie burger, are analog in general appearance but not taste or texture. Non-analog PBMAAs are based on traditional non-meat dishes and have not been altered to look or feel like meat but can be used in the same consumption contexts.

Beyond examination of different PBMA types, research is only beginning to understand

the role prices play in PBMA preference and choice [16, 17]. In the U.S., for instance, prices for beef alternatives exceed those for beef by 20% [18]. In a study among 1,039 German adult consumers, both omnivores and flexitarians mentioned that they perceive meat to perform better than meat alternatives in terms of price [19]. Given the high price tag of some of these products, there has been an ongoing discussion on the need to make PBMA more affordable. Fast-food giant McDonald’s, for instance, withdrew its meat-mimicking McPlant burger from the U.S. market after poor sales across 600 test restaurants [20]. Customer feedback indicates that, despite its appealing taste, the McPlant was considered too expensive: *“It tastes like real meat, but it’s too expensive to buy again”* [21]. Price may partly explain why studies find PBMA market shares to be limited in the 20-25% range [16, 17, 22]. In response, grocery chain Lidl in Germany launched an ambitious initiative to achieve price parity between PBMA and their meat counterparts [23]. However, it raises a pivotal question: can achieving price parity alone instigate a significant shift in protein consumption, or is it imperative for PBMA to surpass meat in affordability? Because plant-based protein has lower per-gram costs than meat [7], greater affordability appears an attainable scenario as soon as processing costs have decreased.

To tackle these research gaps, we conducted two large-scale studies—one survey ($N_{\text{STUDY1}} = 1,003$) and one experiment ($N_{\text{STUDY2}} = 1,123$)—to examine preferences for certain PBMA. In the survey, we considered rankings of a meat burger and three PBMA burgers as well as genuine consideration of these options to determine the heterogenous preference distribution. In the experiment, we tested the impact of price on burger consideration and choice.

Study 1: Preference Heterogeneity

Aims and Design. The goal of this study was to gain a better understanding of preferences for various PBMA. We showed 1,003 American adults images and patty ingredient lists of four burger alternatives: meat (beef burger), analog (which mimics meat; plant-based burger), semi-analog (analog in general appearance but not taste or texture; veggie burger),

and non-analog (falafel burger) (*Methods* and *SI Appendix*, Fig. S1). We asked respondents to rank all four burgers and to indicate, for each burger option, if they would genuinely consider purchasing it outside the study context.

Results

Preferences: model-free evidence. The ranking task reveals that, unsurprisingly, the meat burger is by far the most popular option (75.0% of first-place votes; *SI Appendix*, Fig. S2). The burger with the second-highest share of first-place votes (11.5%), as well as the largest share of second-place votes (33.7%), is the non-analog falafel burger. The semi-analog and analog—both options that mimic meat, albeit to different degrees—are rated very similarly (7.3% and 6.3% of first-place votes, respectively). The largest share of “last-place votes” is received by the meat analog. Consideration is highest for the meat burger (91%) but is also sizeable for PBMA, with shares ranging between 48.5% (semi-analog burger) and 43.3% (analog burger). Burger options that rank first are almost always considered (99%), whereas lower ranks have lower consideration (69.4%, 42.3%, and 17.1% for the ranks 2 to 4, respectively). This makes intuitive sense and highlights the value of the data for our analysis, as it provides valuable information about burger preferences beyond first choices.

Preferences: model-based evidence. We estimate a hierarchical exploded logit model [24, 25] to incorporate multiple-ranked choices for each person, consideration (in addition to the ranking for anchoring purposes [26]), as well as unobserved and observed heterogeneity (*Methods* and *SI Appendix*, *SI Text*). Overall, we find that the meat burger has higher utility than all PBMA (*SI Appendix*, Table S2). Results further point to preference differences regarding the observed heterogeneity (Fig. 1). For example, female respondents have lower utility for the meat burger but higher utility for the semi-analog burger (compared to males). Similarly, high education is associated with lower meat burger utility but higher analog and non-analog burger utility (compared to low education). The non-analog burger has shrinking utility with increasing age (by 10-year cohort).

>>> Insert Fig. 1 about here <<<

After controlling for observed heterogeneity, we still find significant σ values for all four burgers (SI Appendix, Table S3), which points to substantial unobserved heterogeneity in burger preferences. Unobserved heterogeneity is largest for the semi-analog burger, indicating that some respondents seem to really prefer the veggie burger, while others strongly dislike it. The magnitudes of the σ values help explain why not all respondents rank the burgers in the same way. Further, for the unobserved heterogeneity, we also estimate correlations between the options' utilities (*SI Appendix*, Table S4) and find that the meat and semi-analog burgers have a significantly negative relationship ($\omega = -0.39$, 95% CI: $[-0.51, -0.26]$). We observe the same relationship between meat and the non-analog burger ($\omega = -0.24$, 95% CI: $[-0.43, 0.14]$). There is a non-significant correlation between the meat and analog burgers ($\omega = 0. - 16$, 95% CI: $[-0.45, 0.15]$), meaning there is no clear relationship in preference between these two options. The correlations between the PBMA options are all positive and significant, pointing to a complementary relationship among the three PBMA types.

Counterfactual simulation. Based on the hierarchical exploded logit model it is possible to run counterfactual (“what if”) simulations [27]. We simulate the market shares for several scenarios with varying availability of burger options while accounting for the estimated preference heterogeneity and consideration effects (*SI Appendix*, SI Text). In a scenario with all four options available (scenario 8 in *SI Appendix*, Table S5), 75.4% of consumers would choose the meat burger, followed by 11.1% choosing the semi-analog, 8.3% the non-analog, and 4.8% the analog burger. We note that these shares align well with the model-free results for the 1st rank shares reported (*SI Appendix*, Fig. S2), lending credibility to the simulation results. The finding suggests that most consumers prefer the original (i.e., meat burger) over the replica (i.e., analog burger) but also that “traditional” PBMA, such as the semi-analog veggie burger and the non-analog falafel burger, are no less popular than meat analogs. Results further indicate that few respondents choose the none option, meaning that collective demand for the 3 PBMA is close to 25%. We note that this estimate is remarkably similar

to that found in previous studies [16, 17, 28]. In contrast to a scenario featuring meat alongside a single PBMA, the introduction of a second and third PBMA indicates a 42.5% and 17.7% overall increase in market shares for PBMA. This implies a notable rise in PBMA share (and a decrease in meat share by 5.6% and 4%, respectively) as the choice set expands to include more PBMA.

A simulation in which the meat option is not available (scenario 15 in *SI Appendix*, Table S5) suggests that market shares would notably increase for each PBMA: analog 15.0% (+212%), semi-analog 27.3% (+146%), and non-analog 26.6% (+220%). Apparently, all PBMA benefit from the unavailability of meat in nearly the same way, including the semi-analog and non-analog burgers. Although 31.2% would leave the market and choose none of the plant-based proteins, collective PBMA demand would increase to 68.8%. This indicates that, with some incentive, consumers might become more open to PBMA, almost tripling market shares relative to the baseline scenario.

Study 2: Price Response

Aims and Design. In this discrete choice experiment, we used a 2 (relative price: 25% higher vs. lower than the baseline price of \$10) \times 4 (burger option) plus control (all prices are at the baseline) between-subjects design to test the impact of price on PBMA consideration as well as choice, aiming to capture both stages of the decision-making process. We aimed for 125 participants per cell due to the anticipated high share of meat burger choice (final $N_{\text{STUDY2}} = 1,123$). We showed participants the same burger options (i.e., meat, analog, semi-analog, and non-analog) as in study 1 and asked them to indicate which they would consider ordering for lunch. Based on their answer, participants were asked to indicate which of the considered burgers they preferred the most.

Results

Substitution and price response: model-free evidence. A descriptive analysis of burger consideration frequencies reveals that most consumers consider the meat burger (85.2%), followed by the non-analog (36.1%), analog (35.8%), and semi-analog burgers (34.8%). In terms of absolute and relative magnitude, these values are similar to those in study 1. A comparison of the empirical distribution of consideration sets and their implied distribution given the aforementioned burger consideration frequencies assuming independence reveals significantly different results (*SI Appendix*, Fig. S4). The discrepancy indicates that accounting for dependencies (negative or positive) between burger consideration is crucial for understanding consumer preferences. The most common consideration set contains only the meat burger (38.4%). Further, consumers tend to consider either one specific PBMA option or all of them, while consideration sets with two plant-based burgers are rare. Similar to what we find in study 1, approximately three-quarters (72.7%) would choose the meat burger, followed by the non-analog (8.3%), analog (8.3%), and semi-analog (6.9%) burgers. In terms of price response, we find that the average consideration for a burger increases by about 9 percentage points at the low price point of \$7.50 compared to the high price point of \$12.50. Similarly, choice increases by about 5 percentage points when comparing burgers offered at \$7.50 versus \$12.50. These differences are economically meaningful.

Substitution: model-based evidence. To combine the consideration and choice stages, we use a two-stage model [29] that first employs a multivariate logit model [30] to analyze burger consideration, allowing for associations between alternatives. In the second stage, we use a multinomial logit model [25] for burger choice, using only the observed consideration sets. In both stages, we account for price effects and observed heterogeneity. Both models fit the data very well, with predicted consideration (choice) shares of 85.1% (72.6%), 36.7% (7.9%), 35.1% (6.2%), and 36.4% (8.1%), for the meat, analog, semi-analog and non-analog burgers, respectively (the predicted consideration shares also align with the data;

SI Appendix, Fig. S4). Parameter estimates (*SI Appendix*, Table S8) indicate that female consumers consider the semi-analog burger more than male consumers ($\theta = 0.69$, 95% CI: [0.38, 0.99]) but consider the analog ($\theta = -0.34$, 95% CI: [-0.62, -0.08]) and non-analog burgers ($\theta = -0.60$, 95% CI: [-0.90, -0.30]) less. Highly educated consumers show no differences in PBMA burger consideration but consider the meat burger less than consumers without a college degree ($\theta = -0.72$, 95% CI: [-1.16, -0.28]). We further observe an uphill battle for PBMA consideration among consumers who self-report to never buy PBMA. These consumers have higher meat consideration ($\theta = 1.54$, 95% CI: [0.75, 2.44]) and lower plant-based burger consideration, irrespective of PBMA type (θ s between -1.41 and -0.63, P s < 0.05). Similar to study 1, we find that meat burgers have the highest utility at the choice stage but differences to the PBMA are less pronounced as we model burger choice conditional on consideration. Parameter estimates (Table S8) show that female consumers have a lower utility for the meat ($\theta = -0.86$, 95% CI: [-1.59, -0.15]) and analog burgers ($\theta = -0.85$, 95% CI: [-1.71, -0.03]). As with the consideration stage, consumers with no prior PMBA experience have a lower utility for PBMA burgers (θ s between -1.64 and -1.22, statistically significant for analog and non-analog burgers).

Echoing substitution patterns from study 1 as well as the model-free results, association patterns between the four burger options indicate that consideration sets that include the meat option may not contain any PBMA (Table 1, lower-left part). For example, individuals are less likely to consider both the meat and non-analog burgers ($\psi = -0.82$, 95% CI: [-1.28, -0.37]). Conversely, consideration of any PBMA is associated with increased consideration of the remaining PBMA. Positive and significant association parameters are found for the semi-analog and analog ($\psi = 0.86$, 95% CI: [0.56, 1.16]) as well as non-analog burger ($\psi = 1.45$, 95% CI: [1.15, 1.77]), respectively. The association between the analog and non-analog burger is non-significant ($\psi = 0.22$, 95% CI: [-0.08, 0.51]). Apparently, individuals seem to consider these specific PBMA independently of each other.

>>> Insert Table 1 about here <<<

Price response: model-based evidence. Price has a positive but non-significant effect on meat consideration ($\theta = 0.17$, 95% CI: $[-0.08, 0.42]$) but a negative and significant effect on PBMA consideration ($\theta = -0.11$, 95% CI: $[-0.18, -0.04]$) and choice for all options ($\theta = -0.19$, 95% CI: $[-0.34, -0.04]$) (*SI Appendix*, Table S8). Being older ($\theta = 0.08$, 95% CI: $[0.01, 0.15]$), identifying as female ($\theta = 0.29$, 95% CI: $[0.05, 0.52]$), and having a college degree ($\theta = 0.35$, 95% CI: $[0.13, 0.58]$) significantly decrease price sensitivity regarding burger choices. To interpret the price effect, we examine price elasticities at the consideration stage as well as total elasticities (i.e., consideration and choice). Total elasticities (Table 2, bottom panel) show how choices of the four burger options plus the option to choose none change depending on price. Own elasticities are negative for all burgers and stronger in absolute terms for PBMA. Specifically, the elasticity for the analog burger is significant and greater than $|-1|$ ($\varepsilon = -1.4$), indicating price-elastic demand. Reducing the analog burger’s price by 10%, for instance, would result in a 14% sales increase. For the semi-analog and non-analog burgers, demand is price-inelastic ($-1 < \varepsilon < 0$). Hence, for the meat burger, we do not observe a meaningful effect of the price on choices. By contrast, for any 1-percent decrease (increase) of the sales price of a given PBMA, the choice of that alternative increases (decreases) by approximately 0.5 to 1.4 percent. We find that choice elasticities are comparable to those in previous studies, especially when considering our out-of-home food demand context [16, 31, 32].

>>> Insert Table 2 about here <<<

Cross-price elasticities indicate that after a PBMA price decrease, the increased PBMA choice comes from individuals who would not have chosen any burger option (i.e., from the outside good) but also from those who choose the meat option when prices are the same across all products. The effect, both in terms of pattern and effect size, is similar for each of the 3 PBMA, but the analog burger has the highest cross-price elasticity with regard to the meat burger ($\varepsilon = 0.13$, 95% CI: $[0.06, 0.20]$). All cross-price elasticities of the meat burger are non-significant, but there is an 87.1% probability that increasing the price of the meat

burger would increase the choice of the analog burger ($\varepsilon = 0.39$, 95% CI: $[-0.27, 1.10]$).

In addition to the pattern described above, significant cross-price elasticities among the PBMA exist at the consideration stage, reflecting the patterns of the association parameters. For example, reducing the price of the semi-analog burger not only increases the likelihood of considering that semi-analog burger ($\varepsilon = -0.56$); it also increases the likelihood of considering the analog ($\varepsilon = -0.12$) as well as non-analog burgers ($\varepsilon = -0.18$). At the same time, consideration of the meat burger becomes slightly less likely ($\varepsilon = 0.02$, 95% CI: $[0.00, 0.04]$). Decomposition of the total effect reveals that about half to almost of the total price effect on PBMA choice comes from the price effect on consideration, hence modeling price effects at both stages is crucial. Both effects, increased consideration of PBMA and decreased consideration of the meat option, are prerequisites for a sustained behavior change. To understand the total effects of specific price scenarios and differences across consumer types, we look at the counterfactual simulations in the next step.

Counterfactual simulation. To understand price response across different price scenarios and consumer types, we show meat and PBMA consideration and choice in Fig. 2 (effects for individual PBMA types are shown in *SI Appendix*, Fig. S5-7). We simulate outcomes for different PBMA prices (between \$12.50 and \$5) and a fixed price for the meat burger (\$10). The counterfactuals underscore the price response such that both meat and non-meat consideration and choice are responsive to price changes, especially large ones. When we compare a scenario of price parity with one in which PBMA cost half the price of meat, meat burger consideration probability drops by 2.5 percentage points (from 85.8% to 83.3%) and meat choice probability by 17 percentage points (from 74% to 57%), while PBMA consideration increases by 13 percentage points (from 61.2% to 74.4%) and choice probability by 16.5 percentage points (from 21.3% to 37.8%). The positive price effect on PBMA consideration is not only visible as an increased likelihood of considering at least 1 PBMA option; we also observe that consideration set size increases (in the baseline condition, only 27.6% of consideration sets include more than 2 burgers; that number increases to 46.6

when the PBMA price decreases).

>>> Insert Fig. 2 about here <<<

Further, individual differences point to consumer types that vary in price response (*SI Appendix*, Table S8). For example, highly educated consumers are generally less likely to consider meat and more likely to consider PBMA (compared to consumers without a college degree, see Fig. 2). As PBMA prices become more competitive, however, the gap widens with regard to PBMA consideration but narrows for PBMA choice. Although U.S. consumers without a college degree consider PBMA less often, lower PBMA prices significantly increase their likelihood of purchasing meat alternatives, narrowing the gap with highly educated consumers.

In terms of gender, males are more likely to consider meat but also PBMA (a look at the individual PBMA types reveals that males are more likely to consider the analog and non-analog burger but less likely to consider the semi-analog burger; *SI Appendix*, Fig. S5). Despite this general difference in burger consideration, males and females respond similarly to price changes. When it comes to burger choice, however, price affects gender very differently. While males become more likely to choose PBMA as their price becomes more competitive, a favorable price has a minimal effect on PBMA choice among females (irrespective of PBMA type).

Counterfactuals also illuminate the type of “no prior PBMA eaters.” Individuals who do not normally eat PBMA seem less normative about it (unlike individuals who state they never eat meat; *SI Appendix*, Fig. S5). This indicates that unfamiliarity with PBMA does not carry forward; instead, there is an openness to at least consider PBMA, especially when they become more affordable (Fig. 2). Stimulating PBMA sales in this type remains an uphill battle though. By contrast, individuals with at least some experience of purchasing PBMA are more willing to refrain from purchasing the meat burger and try PBMA instead, if the price becomes attractive. When we simulate education and gender effects only within the group of individuals with at least some PBMA experience, we find that males without

a college degree would be almost indifferent between a meat and analog burger (each choice probability $\sim 30\%$) when the latter costs half the price (*SI Appendix*, Fig. S7).

Discussion

Consumers overwhelmingly prefer meat over PBMA, yet the landscape becomes less clear when determining which alternatives to meat they might also consider and the conditions under which shifts in preference occur. This study sought to untangle these complexities, providing a deeper understanding of consumer preferences in this domain. Our findings underscore the necessity of differentiating between various categories of PBMA, given the increasing diversity of options on the market, rather than treating them as a homogeneous group, a distinction often overlooked in existing studies [6]. This nuanced approach becomes crucial when examining preferences and substitution patterns.

First, PBMA, particularly when various types like analog, semi-analog, and non-analog are offered together, are appealing enough to be selected in two out of three cases when no meat option is available. Interestingly, among the three types of PBMA, no unanimous picture of consumer preferences emerged. The semi-analog and non-analog burgers would have similar market shares if offered exclusively but also if offered together. The analog burger was less popular but not far behind the other two. The finding can help explain why some research indicates that consumers prefer alternatives that closely resemble meat in flavor, texture, appearance, and smell [33], whereas others find that consumers prefer lightly processed options that do not resemble meat over meat replacements [14]. Indeed, PBMA appear to share commonalities that allow spillover effects. For example, consideration and preference for one PBMA type increases consideration and preference for others.

Second, the availability of a meat option dramatically shifts choice, with three-quarters of individuals opting for meat over PBMA. We find heightened PBMA preference among consumers with a college degree as well as female consumers who prefer the semi-analog burger. By contrast, consumers who self-report never eating PBMA find extremely little

utility in eating PBMA, including those that mimic meat. Yet we observe that a majority of consumers would at least consider a PBMA option alongside meat even if they ultimately choose meat. One may speculate that a person who chooses meat after considering a PBMA, might opt for the PBMA on a subsequent occasion. Additionally, consumers who consider both meat and non-meat options may be more open to hybrid products that combine meat and plant-based ingredients. Supporting this, a study among 99 UK consumers found that after trying beef, plant-based, and hybrid burgers, acceptance for the hybrid burger was greater than that of the plant-based burger in both blind and informed conditions [34].

Third, a potentially unintended consequence of spillover effects is partial substitution among PBMA. In study 1’s market share simulations (*SI Appendix*, Table S5), we observe that the market share for each PBMA type would increase by at least 5 percentage points, if the other PBMA were not available. We further observe diminishing returns of adding more PBMA variety, which could prevent these options from becoming profitable. These findings corroborate the concerns of a growing number of scholars [32, 33, 34, 35]. While many assume that stimulating consumption of PBMA will lead to a reduction in meat consumption, this substitution pattern indicates that these products are not currently fulfilling their sustainability goal of displacing meat on a large scale. Still, as we introduce PBMA into the choice set (study 1), there is a discernible shift in overall market share, with an increase in the PBMA share and a corresponding decrease in the meat share. Further, reducing PBMA prices “steals” meat choices but does not harm the choice of the remaining PBMA (study 2). While our study’s methodology and design differ, this trend aligns with the observations made by Garnett and colleagues, where increasing vegetarian options in a university dining hall boosted vegetarian sales and reduced meat sales [35].

A fourth finding is that prices affect demand for PBMA. When we factor in pricing, the preference for PBMA falls below 20% if they are priced higher than meat, which is typically the case. If PBMA and meat are priced equally, the preference for PBMA increases to 21%, mirroring findings from study 1 and previous research [16, 17, 28]. A further reduction in PBMA prices can significantly boost their popularity. For instance, if PBMA are priced

at half the price of meat, their choice shares increase to 38%, and nearly 50% among male consumers. However, it is important to note that these figures are based on a combined offering of various PBMA types. For restaurants and producers of PBMA, price-elastic demand means that revenues would actually increase when they offer different PBMA types at lower prices than the meat options. Importantly, we show that price variation does not only impact choice but also consideration set size and composition. This means that even if initiatives to promote PBMA consumption may not show desired outcomes right away, the typically unnoted benefit of increased consideration could be accomplished.

Our studies have several strengths. We use two large samples representative of the United States to explore preference patterns for multiple PBMA along the consumer consideration and choice stages. Notably, our unique contribution lies not only in the subject matter but also in the innovative methodology employed. In both studies, we were concerned with complex substitution patterns, particularly between meat and PBMA burgers. In study 1, we addressed this with a flexible specification for correlated unobserved heterogeneity. To obtain more information from each respondent, we opted for burger ranks as input for our exploded logit model. In study 2, we explicitly added burger prices as drivers for both decision stages, consideration and choice. We refrained from using the standard approach in choice-based conjoint (CBC) analysis that asks consumers to make repeated choices. Although within-subject designs allow the researcher to model unobserved preference heterogeneity, their predictive validity can drop as participants utilize decision heuristics [36, 37]. As we did not want to measure price effects within respondents but in a purely between-subjects design, we opted for a simpler multivariate logit model at the consideration stage that still can account for interdependencies between alternatives without needing multiple observations per respondent. As our results show, modeling consideration and choice separately is crucial, as prices affect both decision stages, resulting in complex elasticity patterns that simple models assuming full consideration at the choice stage would not be able to infer. Similar results can be found in an online shopping context [29]. Lastly, working with probability models and Bayesian estimation methods allows us to easily perform simulations that include the

full uncertainty from the estimation.

However, our studies also have limitations that provide opportunities for future research. To gain an in-depth understanding of substitution patterns and price response, we focused on a single product category (burgers) and large samples from a single country (United States) under hypothetical conditions. While this decision benefits internal validity, we are careful to claim generalizability to other product categories and countries. This further includes the visual presentation of our stimuli, where AI was used to create images of an analog burger that convincingly looked like a meat burger. However, commercial products have struggled to achieve such visual parity with animal-based burgers. Thus, our estimates of preferences for meat analogs could be on the optimistic end of the spectrum. Future research could investigate a more diverse set of countries and products, including novel and ostensibly more sustainable animal-based proteins such as cultured meat. Ideally, these studies would be able to incorporate revealed preferences in addition to using hypothetical choice. A CBC analysis with repeated measurement that accounts for rich unobserved heterogeneity (with regard to consideration sets and preferences) would allow for flexibly estimating individual-level price parameters for targeted promotions and product recommendations. Such an application has practical relevance for online food delivery services that have detailed information about their customers and can reach them at the individual level via apps.

Future research might also explore the diverse nutritional profiles among PBMA, which correspond to the variety of product offerings. Traditional vegetarian alternatives are generally made from whole legumes and grains, offering a nutrient-dense, low in saturated fat option with health benefits. Others, like meat analogs, can be categorized as ultra-processed foods (though proponents argue they differ from other foods in this category, such as soda and confections [38]). While evidence on the health value of meat analogs and their ability to replicate the nutritional profile of meat equivalents is limited, some studies suggest these foods can be associated with positive health outcomes, and processing whole-plant foods into protein isolates may not necessarily compromise their health value. Processing can improve product safety and enable fortification and enrichment [39]. For instance, the Nutrition Facts

Label on many of these meat-mimicking products matches real meat in levels of protein and B12, offering a comparable nutritional profile.

Another fruitful area for future research is how the marketing of PBMA can increase preference beyond price effects. An interesting conversation in this regard is about which benefits of PBMA consumption should be highlighted to make them more competitive vis-a-vis meat. Some researchers argue that the health and sustainability advantages should be highlighted, because this is the key relative strength of PBMA [40]. Other researchers, however, argue that the good taste of PBMA should be highlighted, because many consumers expect PBMA to be less tasty than meat [41, 42]. Against this background, it has been found that advertised benefits should match an active eating goal, such as promoting sustainability when a sustainability goal is active but taste when a hedonic goal is active [43].

Implications

In the face of the pressing need to curtail animal consumption for environmental sustainability, our research investigates the complexities of the plant-based meat alternative (PBMA) market, particularly in the United States, one of the highest beef-consuming countries globally, with a per capita consumption of 25.32 kilograms in 2023 [3]. Our studies uncover significant heterogeneity in the U.S. market for PBMA, with varying personal preferences and types showcasing diverse tastes. A substantial portion of consumers considering meat are reluctant to consider plant-based alternatives, posing a challenge for behavior change (for a similar finding, see [44]). However, leveraging competitive PBMA prices can enhance the likelihood of consideration, translating into greater choice. Importantly, increased consideration for one PBMA extends to other PBMA, emphasizing the interconnected nature of these products. The association between PBMA types, combined with a strong preference for meat, also point to diminishing returns of meat displacement when the variety of available PBMA options would increase. Rather, the outcome would be comparable to micro

targeting, with each PBMA attracting a specific consumer type. While this would make consumers of PBMA happier overall, PBMA availability alone does not seem capable of attaining the goal of a sustainable protein transition. Our findings challenge the notion that PBMA will naturally replace meat, instead aligning with recent findings that meat and meat analogs may indeed complement each other rather than serve as substitutes [13, 31, 45, 46]. However, this dynamic could shift dramatically with more competitive PBMA pricing, potentially turning these products into true meat substitutes. This is particularly true for analogs, which appear to benefit the most from increased affordability.

In light of retail prices posing a challenge for PBMA adoption, and despite their lower per-gram protein production costs compared to meat, our research gives perspectives on future pathways. In part, if these higher prices are related to processing costs [7], there is optimism that as economies of scale are realized, the reduced production cost will equalize the playing field between PBMA and their meat counterparts [47]. Realization of economies of scale may take some time, as PBMA still had a 20% price premium over beef in 2023 [18]. In the long run, the lower production and processing cost of plant proteins could translate to not just equal but lower prices than meat. Nonetheless, bringing PBMA into the consideration set and shopping baskets demands substantial price reductions. While these reductions can be revenue-increasing with elastic demand, restaurants and PBMA producers might be reluctant to reduce prices substantially. At the same time, some countries (e.g., Germany, Netherlands, and Denmark) are contemplating a meat tax, though politically difficult to implement, a policy intervention anticipated to exert significant steering effects on consumers [48]. Some scholars propose allocating tax revenues towards subsidizing plant-based foods, a measure which would further address reluctances of producers [49]. Such measures could heighten the relative economic appeal of PBMA even at moderate price reductions, offering an additional impetus for their widespread adoption. Importantly, however, our findings highlight the importance of affordability beyond price parity in catalyzing the shift towards plant-based diets.

Methods

The research was approved by the Martin Luther University Halle-Wittenberg ethics committee. Participants gave their informed consent at the beginning of each study.

Participants

Across two studies, 2,126 participants (1,196 female, 879 male, 40 non-binary, and 11 declined to answer), aged 18-86 (18-24 = 189; 25-44 = 1,134; 45-64 = 637; Over 65 = 166) were recruited from Prolific ($N_{\text{STUDY1}} = 1,003$, $N_{\text{STUDY2}} = 1,123$). Recruitment was limited to the United States, and participation in study 1 precluded participation in the subsequent study. Across both studies, two participants were excluded for inconsistent response patterns, such as choosing meat while stating never to eat meat, and one participant was excluded for unauthorized repeated participation.

Materials

Each study presented four burger alternatives: meat (beef burger), analog (plant-based burger), semi-analog (veggie burger), or non-analog (falafel burger). We decided to simply name the analog burger “plant-based burger,” which is not uncommon for meat-mimicking burgers. Veggie burgers do not try to replicate the taste and texture of meat, although the visual appearance is similar. Hence, we named the semi-analog burger accordingly. As a non-analog burger, we went with a falafel burger as it is “borrowing” its patty from another popular, traditional vegan dish. Although the falafel burger is not widely available in Western restaurants, fast-food chain Shake Shack offers a falafel burger in the Middle East.

As burger quality varies and inconsistent quality inferences may affect stated preferences across both studies (e.g., comparing a high-quality meat burger with a low-quality non-meat burger, or vice versa), we provided images that looked appealing and stated the typical price point (\$10) for the burgers to signal a medium-to-high quality level (*SI Appendix*, Fig. S1, S3). We pretested burger images with an independent sample through Prolific

($N_{\text{PRETEST}} = 100$). All images were created with Microsoft Designer’s image generator to maintain a comparable appearance. Because we needed the analog burger to look like a meat replica, and the respective AI-generated images did not look like a realistic meat burger, we generated two images of a beef burger, one of which would then be used as the depiction of a meat-mimicking plant-based burger. We pretested both images to examine if they were perceived differently. Because pretest participants evaluated both images the same, we used one image for the analog burger in the main studies and the other image for the meat burger. Below each image, we listed patty ingredients taken from commercially available products.

Pretest participants saw 3 images (semi-analog, non-analog, and one of the two analog burger images) and rated each image on how similar to meat they expected the burger to be in terms of taste, texture, and look. As expected, both versions of the analog burger were perceived as being more similar to meat ($M_{\text{ANALOG}} = 5.34$; $M_{\text{SEMIANALOG}} = 2.44$; $M_{\text{NONANALOG}} = 2.36$; $P_s < 0.001$), while no difference was observed between the two images of the analog burger ($P = 0.60$). Participants also rated how appealing they found each option, with the majority (between 64% and 78%) stating the burgers looked (somewhat) appealing.

Procedure

In both studies, participants were asked to imagine they were at a hamburger restaurant for lunch and that the daily menu listed four burger options that were all the same size. In contrast to previous research [22], we did not instruct participants to imagine that “all burgers taste the same” and “have roughly equivalent nutritional content,” because it is consumers’ varying perceptions about these issues that shape their preferences. For example, there is robust evidence that consumers expect meat burgers to be tastier than non-meat burgers [28, 50]. As we examined price effects in study 2, we added a sentence to provide a price anchor (“Typically, burgers cost around \$10 (fries included)”).

Study 1. After the description of the setting, respondents were first asked which option they preferred for lunch. On a new page, they were asked to imagine the option was already sold out and state their second preference (we displayed the remaining three options). On yet another page, we asked respondents to indicate which of the remaining two options they preferred to get a full ranking of all options sequentially. After the ranking, we asked respondents to indicate, for each burger option, if they would genuinely consider purchasing it. After these tasks, we collected demographic information.

Study 2. Participants were randomly assigned to 1 of 9 conditions as part of a 2 (relative price: 25% higher vs. lower than the baseline price) \times 4 (burger option) plus baseline condition between-subjects design. In the baseline condition, all four burgers cost \$10. In the remaining conditions, the price was changed to either \$12.50 (higher relative price) or \$7.50 (lower relative price) for one burger at a time, while the remaining three burgers had a price of \$10 (*SI Appendix*, Table S6 lists burger prices at selected hamburger restaurant chains). To increase the realism of the study, we presented the burgers in ways similar to restaurant menus with a short display of key burger ingredients, rather than a full ingredient list of the patty. Because the ingredients were not provided, naming one option the “plant-based burger” may have irritated participants making them wonder if the veggie and falafel burgers may not be plant-based. In response, we named the meat analog burger as a burger with a “tastes-like-meat” patty.

Across conditions, participants were first asked to select any and all burgers that they could imagine ordering for lunch. In a funneled presentation, participants were then asked to choose which of the previously selected options they preferred the most. Participants subsequently stated their purchase likelihood for the respective choice on a 7-point rating scale. After these tasks, we collected demographic information.

Weighting

To gain samples representative for the adult U.S. population, we weighted responses based on gender identity, age, and education (*SI Appendix*, Tables S1 and S7). We note that results did not substantially vary from an analysis without weights..

Analyses

We provide an extended technical appendix in the *SI Appendix*. All analyses were performed in Stan and R.

Study 1. We estimate a hierarchical exploded logit model [24, 25] to incorporate multiple-ranked choices for each person (not just the first choice), consideration, and individual heterogeneity. Based on this model, it is possible to infer realistic counterfactuals that can be used to simulate market shares across “what if” scenarios.

Study 2. In a two-stage model, we first estimate a multivariate logit model [30] to capture the interdependencies among burger alternatives, the influence of prices, and the effect of observed heterogeneity at the consideration stage and then multinomial logit model for burger choice [25], conditioned on the observed consideration sets (i.e., for each respondent, only the considered burgers enter the choice model). Combining both stages, the final (unconditional) burger choice probability is calculated by multiplying the consideration set probabilities with the conditional choice probabilities [29, 51]. As we explicitly collect the consideration set and choice information, both stages are easy to model and estimate. Note that both stages do not share parameters and can be separately estimated [29] to accommodate flexible substitution patterns.

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References

- [1] Parlasca M, Qaim M (2022) Meat consumption and sustainability. *Annual Review of Resource Economics* 14:17–41.
- [2] Biesbroek S, et al. (2023) Toward healthy and sustainable diets for the 21st century: Importance of sociocultural and economic considerations. *Proceedings of the National Academy of Sciences* 120(26):e2219272120.
- [3] OECD (2023) Meat consumption (indicator). Available at: <https://doi.org/10.1787/fa290fd0-en> [Accessed November 28, 2023].
- [4] Onwezen M, Bouwman E, Reinders M, Dagevos H (2021) A systematic review on consumer acceptance of alternative proteins: Pulses, algae, insects, plant-based meat alternatives, and cultured meat. *Appetite* 159:105058.
- [5] Circus V, Robison R (2019) Exploring perceptions of sustainable proteins and meat attachment. *British Food Journal* 121(2):533–545.
- [6] Fischer AR, Onwezen MC, van der Meer M (2023) Consumer perceptions of different protein alternatives in *Meat and Meat Replacements: An Interdisciplinary Assessment of Current Status and Future Direction*, eds. Meiselman HL, Manuel Lorenzo J. (Elsevier), pp. 333–362.
- [7] Rubio N, Xiang N, Kaplan D (2020) Plant-based and cell-based approaches to meat production. *Nature Communications* 11:6276.
- [8] He J, Evans N, Liu H, Shao S (2020) A review of research on plant-based meat alternatives: Driving forces, history, manufacturing, and consumer attitudes. *Comprehensive Reviews in Food Science and Food Safety* 19:2639–2656.
- [9] Jahn S, Furchheim P, Strässner A (2021) Plant-based meat alternatives: Motivational adoption barriers and solutions. *Sustainability* 13(24):13271.
- [10] Kumar P, et al. (2017) Meat analogues: Health promising sustainable meat substitutes. *Critical Reviews in Food Science and Nutrition* 57(5):923–932.
- [11] Unilever (2023) New partnership takes plant-based kebab mainstream. Available at: <https://www.unilever.com/news/news-search/2023/new-partnership-takes-plantbased-kebab-mainstream/> [Accessed November 28, 2023].
- [12] Webber J (2023) Beyond Meat sales fall by 30%, reigniting ‘plant-based fad’ debate. Available at: <https://plantbasednews.org/news/economics/vegan-protein-alternatives-market-fad-beyond-meat/> [Accessed November 28, 2023].
- [13] Neuhofer Z, Lusk J (2022) Most plant-based meat alternative buyers also buy meat: an analysis of household demographics, habit formation, and buying behavior among meat alternative buyers. *Scientific Reports* 12(1):13062.

- [14] Rööß E, de Groot A, Stephan A (2022) Meat tastes good, legumes are healthy and meat substitutes are still strange – The practice of protein consumption among Swedish consumers. *Appetite* 174:106002.
- [15] Web of Science (2024) Search query. Available at: <https://www.webofscience.com/wos/woscc/summary/140a9085-562f-4214-93a3-958d4656b56c-fd71fda1/relevance/1>. [Accessed July 24, 2024].
- [16] Tonsor G, Lusk J, Schroeder T (2023) Market potential of new plant-based protein alternatives: Insights from four us consumer experiments. *Applied Economic Perspectives and Policy* 45:164–181.
- [17] Van Loo E, Caputo V, Lusk J (2020) Consumer preferences for farm-raised meat, lab-grown meat, and plant-based meat alternatives: Does information or brand matter? *Food Policy* 95:101931.
- [18] Good Food Institute (2024) State of the Industry Report. Plant-based: meat, seafood, eggs, and dairy. Available at: <https://gfi.org/resource/plant-based-meat-eggs-and-dairy-state-of-the-industry-report/> [Accessed July 29, 2024].
- [19] Michel F, Hartmann C, Siegrist M (2021) Consumers’ associations, perceptions and acceptance of meat and plant-based meat alternatives. *Food Quality and Preference* 87:104063.
- [20] Castrodale J (2024) Here’s why McDonald’s scrapped its plant-based burger in America. Available at: <https://www.foodandwine.com/mcdonalds-mcplant-burger-discontinued-8671341> [Accessed July 26, 2024].
- [21] Nieves S (2022) I tried the new McDonald’s McPlant burger. It tastes like real meat, but it’s too expensive to buy again. Available at: <https://www.foodandwine.com/mcdonalds-mcplant-burger-discontinued-8671341> [Accessed July 29, 2024].
- [22] Slade P (2018) If you build it, will they eat it? Consumer preferences for plant-based and cultured meat burgers. *Appetite* 125:428–437.
- [23] Vegconomist (2023) Lidl announces price parity of vegan products with animal based counterparts. Available at: <https://vegconomist.com/retail-e-commerce/lidl-price-parity-vegan-products-animal-counterparts/>.
- [24] Dyachenko T, Reczek R, Allenby G (2014) Models of sequential evaluation in best-worst choice tasks. *Marketing Science* 33(6):828–848.
- [25] Train K (2009) *Discrete Choice Methods with Simulation*. (Cambridge University Press), 2nd edition.
- [26] Schramm JB, Lichters M (2024) Incentive alignment in anchored MaxDiff yields superior predictive validity. *Marketing Letters*. Advance online publication: <https://link.springer.com/article/10.1007/s11002-023-09714-2> [Accessed August 4, 2024].

- [27] Chernozhukov V, Fernandez-Val I, Melly B (2013) Inference on counterfactual distributions. *Econometrica* 81(6):2205–2268.
- [28] Caputo V, Sogari G, Van Loo E (2023) Do plant-based and blend meat alternatives taste like meat? a combined sensory and choice experiment study. *Applied Economic Perspectives and Policy* 45(1):86–105.
- [29] Amano T, Rhodes A, Seiler S (2022) Flexible demand estimation with search data. Available at SSRN: <https://doi.org/10.2139/ssrn.3214812> [Accessed August 4, 2024].
- [30] Besag J (1974) Spatial interaction and the statistical analysis of lattice systems. *Journal of the Royal Statistical Society: Series B (Methodological)* 36(2):192–225.
- [31] Zhao S, Wang L, Hu W, Zheng Y (2023) Meet the meatless: Demand for new generation plant-based meat alternatives. *Applied Economic Perspectives and Policy* 45(1):4–21.
- [32] Law C, Smith R, Cornelsen L (2022) Place matters: Out-of-home demand for food and beverages in great britain. *Food Policy* 107:102215.
- [33] Hoek A, , et al. (2011) Replacement of meat by meat substitutes. a survey on person- and product-related factors in consumer acceptance. *Appetite* 56(3):662–673.
- [34] Grasso S, Rondoni A, Bari R, Smith R, Mansilla N (2022) Effect of information on consumers’ sensory evaluation of beef, plant-based and hybrid beef burgers. *Food Quality and Preference* 96:104417.
- [35] Garnett E, Balmford A, Sandbrook C, Pilling M, Marteau T (2019) Impact of increasing vegetarian availability on meal selection and sales in cafeterias. *Proceedings of the National Academy of Sciences* 116(42):20923–20929.
- [36] Sablotny-Wackershauser V, Lichters M, Guhl D, Bengart P, Vogt B (2024) Crossing incentive alignment and adaptive designs in choice-based conjoint: A fruitful endeavor. *Journal of the Academy of Marketing Science* 52(3):610–633.
- [37] Li Y, et al. (2022) The more you ask, the less you get: When additional questions hurt external validity. *Journal of Marketing Research* 59(5):963–982.
- [38] Chapman J (2024) Processing the discourse over plant-based meat. Available at: https://knowledge4policy.ec.europa.eu/publication/processing-discourse-over-plant-based-meat_en [Accessed July 29, 2024].
- [39] Flint M, Bowles S, Lynn A, Paxman J (2023) Novel plant-based meat alternatives: future opportunities and health considerations. *Proceedings of the Nutrition Society* 82:370–385.
- [40] Ye T, Mattila A (2021) The effect of ad appeals and message framing on consumer responses to plant-based menu items. *International Journal of Hospitality Management* 95:102917.

- [41] Chandon P, Cornil Y (2022) More value from less food? effects of epicurean labeling on moderate eating in the united states and in france. *Appetite* 178:106262.
- [42] Turnwald B, Crum A (2019) Smart food policy for healthy food labeling: Leading with taste, not healthiness, to shift consumption and enjoyment of healthy foods. *Preventive Medicine* 119:7–13.
- [43] Erhard A, Jahn S, Boztug Y (2024) Tasty or sustainable? goal conflict in plant-based food choice. *Food Quality and Preference* 120:105237.
- [44] Lemken D, Spiller A, Schulze-Ehlers B (2019) More room for legume – consumer acceptance of meat substitution with classic, processed and meat-resembling legume products. *Appetite* 143:104412.
- [45] Siegrist M, Hartmann C (2023) Why alternative proteins will not disrupt the meat industry. *Meat Science* 203:109223.
- [46] Cuffey J, Chenarides L, Li W, Zhao S (2023) Consumer spending patterns for plant-based meat alternatives. *Applied Economic Perspectives and Policy* 45(1):63–85.
- [47] Chafin C, Larson B (2022) Plant-based protein: parity on the horizon. Available at: <https://www.kearney.com/industry/consumer-retail/article/-/insights/plant-based-protein-parity-on-the-horizon> [Accessed November 28, 2023].
- [48] Roosen J, Staudigel M, Rahbauer S (2022) Demand elasticities for fresh meat and welfare effects of meat taxes in germany. *Food Policy* 106:102194.
- [49] Fesenfeld L, Wicki M, Sun Y, Bernauer T (2020) Policy packaging can make food system transformation feasible. *Nature Food* 1:173–182.
- [50] Schouteten JJ, et al. (2016) Emotional and sensory profiling of insect-, plant- and meat-based burgers under blind, expected and informed conditions. *Food Quality and Preference* 52:27–31.
- [51] Manski C (1977) The structure of random utility models. *Theory and Decision* 8:229–254.

Figures and Tables

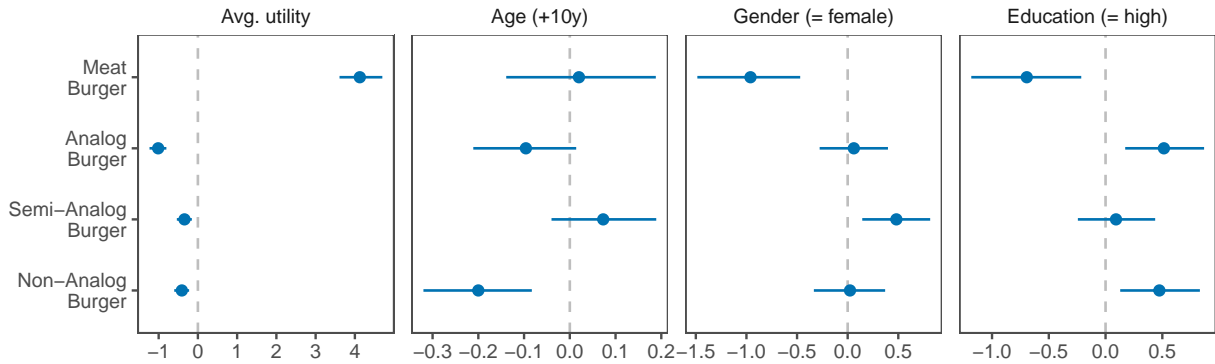


Fig. 1. Exemplary Individual Differences in Burger Preference (Study 1).

Note: The first box shows the average utility of each burger option that accounts for heterogeneity in preferences of the U.S. population. The remaining boxes show observed heterogeneity as planned contrasts (posterior mean plus 95% credible intervals). For example, the meat burger has lower utility for females than males, whereas the semi-analog burger has higher utility for females. The complete results are shown in the *SI Appendix*, Table S2.

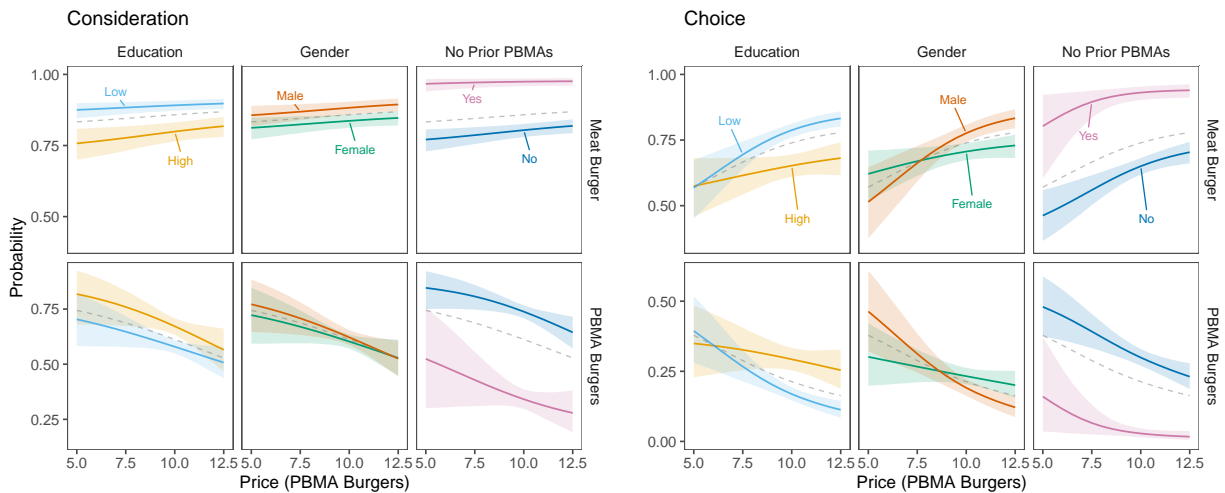


Fig. 2. Price Effect Simulations for Meat and PBMA Burgers (Study 2).

Note: Consideration and choice probabilities are shown for the meat burger and PBMA (aggregated) given a PBMA price range between \$12.50 and \$5, along with 95% credible intervals. Across PBMA price scenarios, the price of the meat burger was kept constant at \$10. A PBMA price of \$10 thus corresponds to price parity. The grey dashed lines display the population mean (representative of the United States). No Prior PBMA's (=Yes) are those who self-report to never eat PBMA's. Disaggregated PBMA results are shown in the *SI Appendix*, Fig. S5-7.

Table 1. Association Parameters and Correlations (Study 2).

Alternative	Meat Burger	Analog Burger	Semi-Analog Burger	Non-Analog Burger
Meat Burger	1	-0.32	-0.55	-0.47
Analog Burger	-0.43 [-0.87, -0.01]	1	0.44	0.31
Semi-Analog Burger	-0.52 [-1.00, -0.05]	0.86 [0.56, 1.16]	1	0.59
Non-Analog Burger	-0.82 [-1.28, -0.37]	0.22 [-0.08, 0.51]	1.45 [1.15, 1.77]	1

Note: Tetrachoric correlations (model-free) are shown in the upper-right part. Association parameters (multivariate logit model) are shown in the lower-left part with 95% credible intervals (CI) in parentheses. The corresponding utility parameters of the multivariate logit model are reported in the *SI Appendix*, Table S8. Significant estimates (at 5%) are bolded.

Table 2. Price Elasticity Decomposition: Consideration and Total Elasticities (Study 2).

Consideration Elasticity								
	Meat Burger		Analog Burger		Semi-Analog Burger		Non-Analog Burger	
	Mean	95% CI	Mean	95% CI	Mean	95% CI	Mean	95% CI
Meat Burger	0.15	[-0.01, 0.31]	0.01	[0.00, 0.03]	0.02	[0.00, 0.04]	0.02	[0.01, 0.04]
Plant-based Burger	-0.05	[-0.14, 0.01]	-0.64	[-1.01, -0.25]	-0.12	[-0.20, -0.04]	-0.07	[-0.14, -0.02]
Semi-Analog Burger	-0.07	[-0.18, 0.01]	-0.12	[-0.21, -0.04]	-0.56	[-0.91, -0.20]	-0.19	[-0.32, -0.07]
Non-Analog Burger	-0.09	[-0.20, 0.01]	-0.07	[-0.14, -0.02]	-0.18	[-0.31, -0.07]	-0.64	[-1.02, -0.26]
Outside Good	-1.46	[-3.05, 0.11]	0.48	[0.15, 0.84]	0.61	[0.20, 1.11]	0.55	[0.19, 0.95]
Total Elasticity								
	Meat Burger		Analog Burger		Semi-Analog Burger		Non-Analog Burger	
	Mean	95% CI	Mean	95% CI	Mean	95% CI	Mean	95% CI
Meat Burger	-0.05	[-0.26, 0.16]	0.13	[0.06, 0.20]	0.08	[0.03, 0.13]	0.05	[0.02, 0.09]
Plant-based Burger	0.39	[-0.27, 1.10]	-1.39	[-2.31, -0.48]	-0.13	[-0.32, 0.05]	0.00	[-0.09, 0.11]
Semi-Analog Burger	0.00	[-0.57, 0.58]	-0.01	[-0.24, 0.23]	-0.46	[-1.42, 0.49]	0.03	[-0.16, 0.26]
Non-Analog Burger	0.01	[-0.56, 0.59]	0.01	[-0.20, 0.20]	-0.32	[-0.65, -0.03]	-0.66	[-1.09, -0.25]
Outside Good	0.09	[-0.89, 1.13]	0.21	[0.05, 0.38]	0.11	[-0.11, 0.34]	0.20	[0.06, 0.38]

Note: Elasticity values are shown with 95% credible intervals (CI) in parentheses. Significant elasticities (at 5%) are bolded. Elasticities represent the percentage change in consideration and choice for a burger alternative in a row in response to a 1-percent price increase of an alternative in a column.

Supporting Information for:
Substitution patterns and price response for
plant-based meat alternatives

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Supporting Information: Text

Study 1. For study 1 we use a hierarchical (or mixed) exploded logit model on the burger rankings (see [1], chapter 7.3 for details). The model is well-suited to our setting as we collected each respondent’s full ranking of all J burger alternatives. It is also easy to incorporate observed and unobserved heterogeneity.

The utility of burger j for respondent i is:

$$u_{ij} = \beta_{ij} + \epsilon_{ij}, \quad \text{where } \epsilon_{ij} \sim EV(0, 1). \quad (1)$$

β_{ij} is the respondent-level burger preference and ϵ_{ij} is the error term. The extreme value distribution assumption of ϵ_{ij} lead to a simple expression (“exploded logit”) for the probability of a specific ranking of the J alternatives based on the utility values (see [2]):

$$Pr(u_{i1} > u_{i2} > \dots > u_{iJ-1} > u_{iJ}) = \prod_{j=1}^J \frac{\exp(\beta_{ij})}{\sum_{k=j}^J \exp(\beta_{ik})}. \quad (2)$$

For the respondent-level burger preferences $\beta_i = [\beta_{i1}, \dots, \beta_{iJ}]'$, we assume a “full” multivariate normal distribution (i.e., a multivariate regression specification, see [3]):

$$\beta_i = \Gamma' x_i + \eta_i, \quad \text{where } \eta_i \sim MVN(0, \Sigma). \quad (3)$$

The matrix Γ includes the parameters to be estimated that relate for each respondent i observed heterogeneity (e.g., demographics) in vector x_i to the burger preferences. A special case of the model without observed heterogeneity would be $\beta_i = b + \eta_i$, where b is simply a vector with population-level average burger preferences.

In addition to the ranking, we ask consideration questions for all alternatives for an “anchoring” of the absolute value of the utilities [4, 5, 6]. To this end we have for each $j \in 1, \dots, J$:

$$v_{ij} = x_{ij} \cdot \beta_i + \nu_{ij}, \quad \text{where } \nu_{ij} \sim EV(0, 1/\mu). \quad (4)$$

The probability of the binary consideration choice (cs) for burger j of respondent i is:

$$Pr(cs_{ij} = 1) = \frac{\exp(\mu \cdot \beta_{ij})}{1 + \exp(\mu \cdot \beta_{ij})}. \quad (5)$$

Hence, we assume that the ranking utilities are the same as the consideration utilities (up to a scale factor μ). The anchor for not considering an alternative is, therefore, 0.

The full likelihood of the model is $l_i^{\text{full}} = (l_i^{\text{rank}} \times l_i^{\text{anchor}})^{w_i}$, and w_i is the normalized survey weight of respondent i . For the hierarchical prior we use a multivariate normal distribution $p(\beta_i) = MVN(b, \Sigma)$, where $\Sigma = \text{diag}(\sigma) \cdot \Omega \cdot \text{diag}(\sigma)$ and $p(b) = N(0, 2.5)$, $p(\sigma) = N_+(0, 1)$. The multivariate normal distribution is common for modeling heterogeneous preferences in choice modeling [1]. The factorization of the covariance matrix Σ into a vector of standard deviations σ and a correlation matrix Ω is advocated in the current literature on Bayesian hierarchical models [7]. The correlation matrix Ω allows the model to capture dependencies between burgers that cannot be fully explained by the observed heterogeneity in x_i . For the correlation matrix we employ the LKJ distribution [8] as prior and shrink the correlation slightly toward zero ($p(\Omega) = LKJ(2)$). This approach has advantages over the common Wishart prior on Σ (e.g., no correlation between components of the covariance matrix; [9]). Lastly we use a lognormal prior on μ to ensure the scale parameter is strictly positive ($\log(\mu) \sim N(0.2, 0.2)$).

There is no available implementation of the model in standard statistical software. Thus, we implemented the model in the probabilistic programming language Stan [10]. Furthermore, we use Hamiltonian Monte-Carlo and the No-U-Turn sampler [11] to draw samples from the posterior distribution. We call our Stan program using the R-package `rstan` in RStudio [12]. Specifically, we run the sampler with five chains for 5,500 iterations with 500 iterations for warmup and keep every fifth draw. This gives us 5,000 draws for each parameter for posterior inference. Chains converge quickly, are stable, and have reasonably high effective sample sizes ($n_{\text{eff}} > 2,500$). All values for the Gelman-Rubin statistic are close to 1 [13]. We also visually inspected trace plots and can confirm convergence and good mixing.

For the counterfactual simulations, we use the posterior draws of β_i and μ . Specifically, we first use the binary consideration probabilities to derive the probability for each possible consideration set. The counterfactual scenarios differ in the available set of alternatives in this step. Next, we compute the conditional choice probabilities from the logit model given each possible consideration set s (i.e., the probability for rank 1 of an alternative). The unconditional choice probabilities follow from weighting the conditional choice probabilities by the corresponding consideration set probabilities ($Pr_{is} = \prod_{j \in s} Pr(cs_{ij} = 1) \prod_{j \notin s} (1 - Pr(cs_{ij} = 1))$). We aggregate over the individuals by using the survey weights to obtain choice shares. Repeating these steps for each posterior draw provides the full posterior distribution of the counterfactual simulation.

Study 2. We use a two-stage model (stage 1: consideration, stage 2: choice) for our discrete choice analysis using a between-subjects design. Two-stage models have a long history in marketing, psychology, and economics (see [14] for an overview of such models using scanner panel data and [15] for state preference data). We were weighing the pros and cons of standard choice-based conjoint (CBC) analysis and our approach concluding the two-stage model is best suited to address our research goal of understanding the effect of prices on burger decisions while explicitly including the consideration stage. As our approach may seem like an unusual design, we will briefly describe our decision-making process:

1. A typical assumption for standard CBC is that all alternatives shown are relevant (i.e., are being considered). If consumers ignore some of the alternatives in the shown choice sets, any inference about preferences under the “full” consideration assumptions will be biased [16]. Specifically, irrelevant alternatives in choice sets lead to an attenuation bias; the impact of attribute changes such as price will be underestimated [15]. Based on findings of study 1, we should expect irrelevant alternatives, such as the meat burger for non-meat eaters or any PBMA burger for “meat lovers” when prices are higher than that of the meat burger.
2. Asking consumers repeatedly about their choices allows modeling unobserved prefer-

ence heterogeneity and, in addition, could provide information about (unobserved) choice set heterogeneity [15, 17, 18]. The potential downside is that many choice tasks are necessary to infer unobserved consideration sets from choices alone (typically more than 10 choice tasks; [15, 17]).

3. Asking many times might then lead to problematic results because respondents pick up simplification strategies and adapt their decision-making process. Li and colleagues show that a CBC study’s external validity drops as early as after 3–6 choice tasks [19]. This even holds for CBC studies with incentive-alignment but less so for adaptive designs (as used by Sawtooth Software) [20]. On the other hand, adaptive designs that maximize statistical efficiency can be cognitively taxing to respondents [19].
4. Another option, compared to a standard CBC setup with an adaptive design and a much more complicated model that allows for unobserved choice set heterogeneity, would be to ask respondents multiple times to state their consideration set in different pricing scenarios (and then continue each time with a choice task conditional on the consideration set). We opted against this because being asked multiple times about product choice is established and common for respondents/consumers as this mimics real life. However, being asked about consideration is unusual. This process tends to involve simple decision heuristics to alleviate the cognitive burden [21], which makes it difficult for consumers to explicitly trace the dynamic consideration set formation across conditions. Consumers are typically not required to explicitly build a consideration set when making purchases (exceptions might be online configurators, etc.). Hence, it might feel strange to respondents having to build choice sets multiple times if we only changed prices for one option at a time. Further, we were concerned about the point above regarding choice task adaptation [19]. If product choices alone are problematic, adding a consideration choice task to each product choice task also adds a lot of effort for respondents.

In line with our research goal of understanding the effect of price on burger decisions

while explicitly including the consideration stage, we chose a practical and straightforward approach. By using a between-subjects design with only one consideration choice and one product choice per respondent, we can efficiently sample a large number of respondents with rich demographic information. We note that this approach explicitly accounts for observed choice set heterogeneity, which we think is the most relevant type of heterogeneity we deal with, reassuring us about the feasibility of our study as it mitigates concerns about unobserved preference heterogeneity.

To model the consideration and choice stages [22], we closely follow Amano and colleagues [23]. In the first stage, we model the consideration of the different burger alternatives using a multivariate logit model [24, 25, 26]. Here, the dependent variable for each respondent is a 4×1 vector that indicates which burger was considered. The multivariate logit model is well-suited for modeling joint discrete decisions, as it allows for 1) interdependencies across alternatives (e.g., due to overlap in features of burgers or correlated preferences), 2) the inclusion of independent variables that may affect the consideration probability for each burger (such as price or demographics), 3) a close form solution for consideration probabilities and the mode likelihood. The use of the multivariate logit model, in contrast to the independent logit model, is a crucial extension of the existing literature [14]. Given our moderately small number of burgers, the total number of possible consideration sets is easily manageable ($2^4 = 16$ possible consideration sets, including the empty set), we use the multinomial logit specification across all consideration sets [25]:

$$Pr(Y_i = y_i | x_i) = \frac{\exp(\mu_{y_i})}{\sum_{s_i \in S} \exp(\mu_{s_i})} \quad (6)$$

with

$$\mu_{y_i} = \sum_{j=1}^J y_{ij} (\alpha_j + x_{ij} \beta_j) + \sum_{k>j} y_{ij} y_{ik} \psi_{jk}. \quad (7)$$

y_i is a specific combination of considered burgers from all possible consideration sets S . x_{ij} is a vector of the respondent and alternative-specific variables (i.e., demographics and price with corresponding interactions). α_j are intercepts for each burger and β_j contain the effect

of x_{ij} on the consideration. The association parameters ψ_{jk} measure whether two burgers are more ($\psi > 0$) or less ($\psi < 0$) likely to be considered together after controlling for the effects of variables in x .

The second stage is a multinomial logit model for the burger choice conditional on the particular consideration set from the first stage. Here the choice probability of respondent i for burger j is:

$$Pr(b_{ij} = 1 | x_i, s_i) = \frac{\exp(u_{ij})}{1 + \sum_{k \in s_i} \exp(u_{ik})} \quad (8)$$

with

$$u_{ij} = \gamma_j + x_{ij}\delta_j. \quad (9)$$

γ_j and δ_j are the intercept and slope parameters of the multinomial logit model. We also ask for purchase intention regarding the most preferred burger option (7-point scale). This allows us to define a cutoff for inferring whether a respondent would buy the burger or opt out. Similar to study 1 this anchoring enables us to estimate the utilities for all options and we restrict the utility for the outside option (i.e., not buying any burger) to 0. We set the cutoff to “5” meaning that we label purchase intention as a purchase if respondents state they were “somewhat likely”, “likely”, or “very likely” to buy the burger. Values below 5 result in an observation for the outside (no purchase) option.

When we combine both stages, we get the final (unconditional) choice probability by multiplying the consideration set probabilities with the conditional choice probabilities. Hence, the unconditional choice probability for a burger j is the weighted sum of the conditional choice probability for a choice set where j is included, with the probability of each relevant choice set as weight. As we explicitly collect the consideration set and choice information, both stages are easy to model and estimate. Note that both stages do not share parameters and could be separately estimated [23]. The likelihood functions for both models are simply the likelihoods of MNL models with alternative- (price) and respondent-specific (e.g., age or gender) variables (see [1] for details). As in study 1, we use normalized survey weights in both likelihood functions. Please note that our model allows for flexible substitution patterns

that do not suffer from the IIA property of the multinomial logit model (independence from irrelevant alternatives [1], as we incorporate discrete consideration set heterogeneity [14]). This is crucial given the product category we investigate, and we know from study 1 that not all burgers are considered by all consumers, and substitution patterns are complex.

As in study 1, we use Bayesian estimation. Combining the intercepts and slope parameters in both models into the vectors $\theta_{\text{consideration}} = (\alpha_1, \dots, \alpha_J, \beta_1, \dots, \beta_J)'$ and $\theta_{\text{choice}} = (\gamma_1, \dots, \gamma_J, \delta_1, \dots, \delta_J)'$ and stacking all association parameters into the vector ψ simplifies the notation. As priors we use $p(\theta_{\text{consideration}}, \psi, \theta_{\text{choice}}) = N(0, 10)$. We also implemented both models in Stan and used for the estimation of each model the same setup as in study 1, to obtain 5,000 draws for posterior inference. As in study 1, visual inspection and formal analysis reveal that the Bayesian estimation worked well ($n_{\text{eff}} > 2,500$ and \hat{R} close to 1).

The procedure for the counterfactual simulations closely follows the description above for study 1. The difference is now that we do not vary the availability of alternatives, but the prices. We simulate consideration set probabilities and unconditional choice probability for PBMA prices between \$5 and \$12.5 (in \$0.5 steps). Again, we use the draws of $\theta_{\text{consideration}}$, ψ , and θ_{choice} . We aggregate over the individuals by using the survey weights to obtain choice shares. Repeating these steps for each posterior draw provides the full posterior distribution of the counterfactual simulation. Note that we also use simulations to obtain arc price elasticities (ε). Note that we also use simulations to obtain arc price elasticities. We simulate a \$0.05 increase and decrease from the baseline price of \$10 for each alternative (including the meat burger) separately and compute the relative changes in consideration and choice probabilities using the midpoint formula.

Supporting Information: Figures

Beef Burger



Patty ingredients:

GROUND BEEF (80% LEAN, 20% FAT), SALT, GARLIC POWDER, ONION POWDER.

Plant-based Burger



Patty ingredients:

WATER, PEA PROTEIN, CANOLA OIL, COCONUT OIL, RICE PROTEIN, FLAVORING, STABILIZER, POTATO STARCH, APPLE EXTRACT, COLOR, MALTODEXTRIN,

POMEGRANATE EXTRACT, SALT, POTASSIUM SALT, CONCENTRATED LEMON JUICE, MAIZE VINEGAR, CARROT POWDER, EMULSIFIER.

Veggie Burger



Patty ingredients:

COOKED QUINOA, RED PEPPERS, COOKED BLACK BEANS, BREAD CRUMBS, EXPELLER PRESSED CANOLA OIL, ROASTED CORN, ONIONS, POTATO FLAKES,

ARROWROOT, TOMATOES, OLIVE OIL, JALAPENO PEPPERS, TOMATO PASTE, EGG WHITE POWDER, ROASTED GARLIC, WHITE VINEGAR, SALT, CILANTRO, CHILI

POWDER, ONION POWDER, CORIANDER, BLACK PEPPER, GARLIC POWDER, CHIPOTLE POWDER, LIME JUICE POWDER.

Falafel Burger



Patty ingredients:

FAVA BEANS, CHICKPEAS, SPICES, SALT, GARLIC POWDER, ONION POWDER, BAKING POWDER, DEHYDRATED PARSLEY, SUNFLOWER OIL.

Fig. S1. Stimuli (Study 1, Presentation Order was Randomized).

Note: The Beef Burger represents the meat option, while the Plant-based Burger represents the analog option. The Veggie Burger and Falafel Burger represent the semi-analog and non-analog options, respectively.

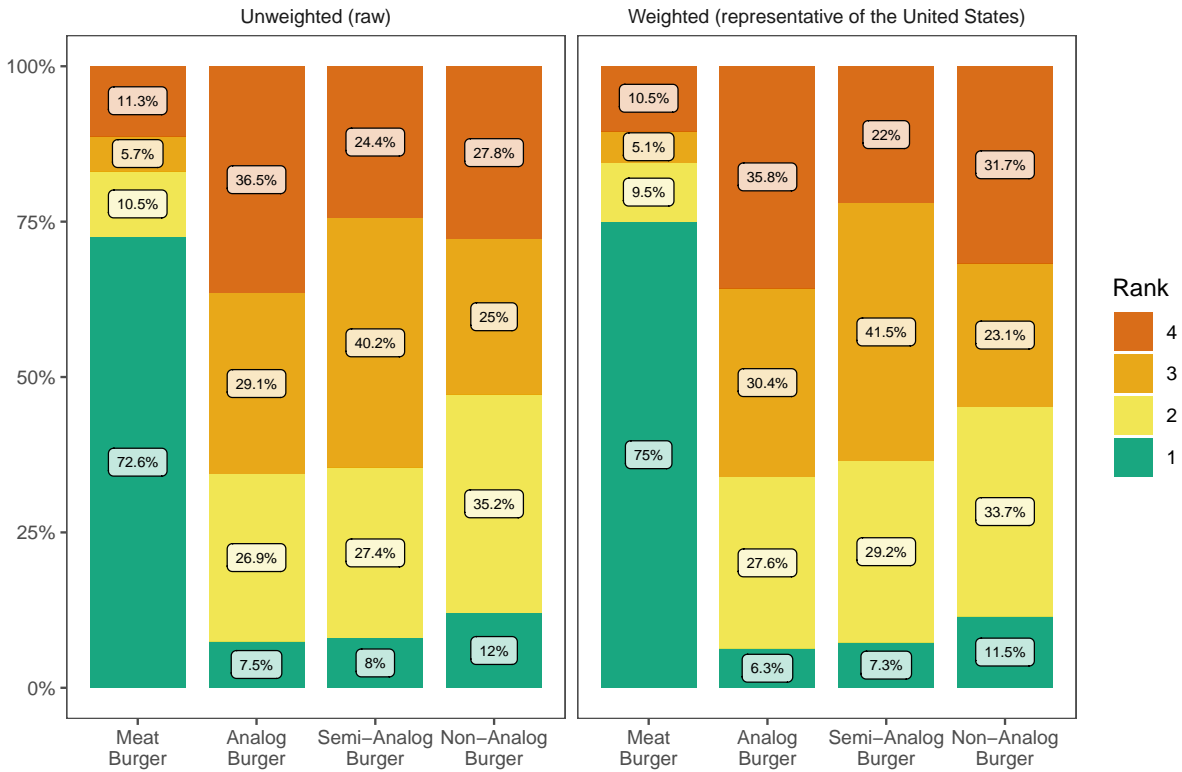


Fig. S2. Ranking of Burger Options (Study 1).

Please select **any and all burgers** that you could imagine ordering in the restaurant. Keep in mind: All burgers are the same size.



Fig. S3. Stimuli (Study 2, Baseline Condition).

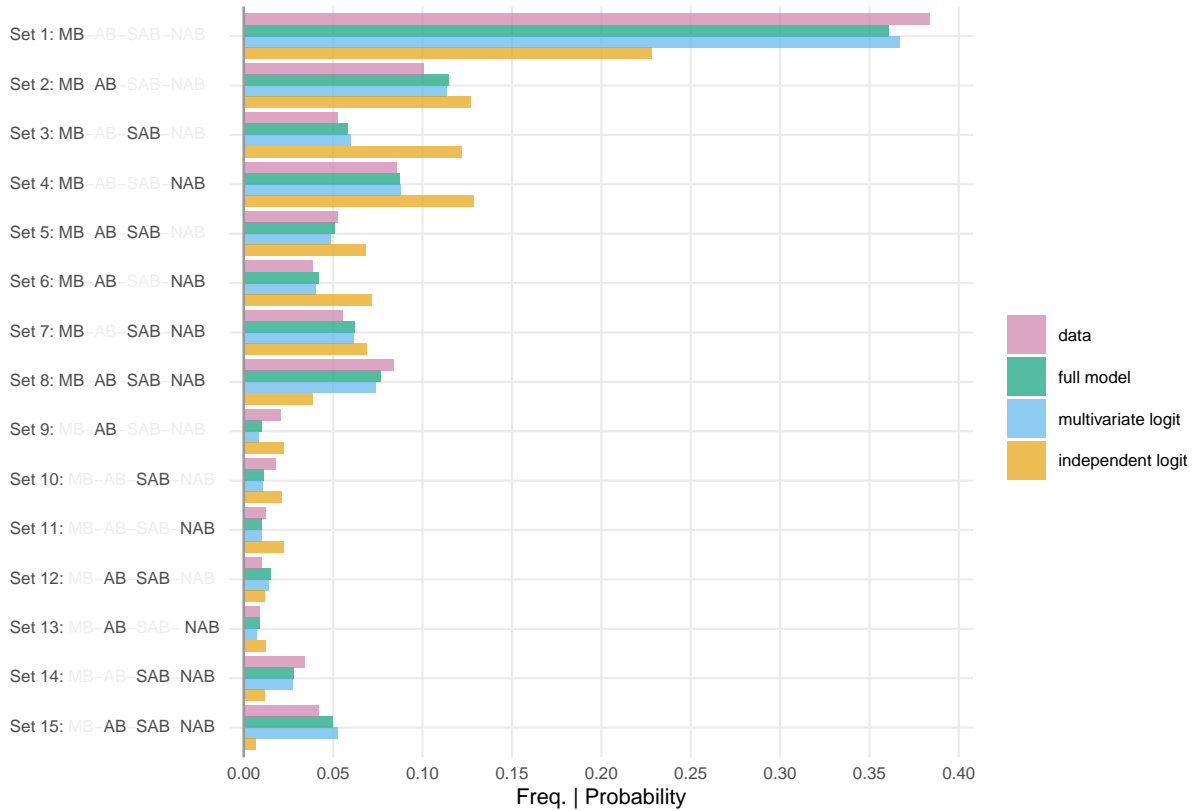


Fig. S4. Consideration Set Frequencies and Model-Based Predictions (Study 2).

Note: The figure shows the empirical frequencies and the model-based results for each possible consideration set as a unique combination of burger alternatives (MB: meat burger, AB: analog burger, SAB: semi-analog burger, NAB: non-analog burger). The independent logit with $Pr_s = \prod_{j \in s} Pr_j \prod_{j \notin s} (1 - Pr_j)$ maps the burger probabilities (i.e., observed frequencies) to consideration set probabilities, assuming independence between alternatives. The figure shows that this model is not suited for our data as the implied consideration set probabilities and the data are significantly different ($\chi^2 = 168.1$, $df = 14$, $P < 0.01$). The data clearly shows interdependencies between burgers as, e.g., empirical frequencies for set 1 are much higher than under the independence assumption or consideration sets including the meat burger and a PBMA burger (set 2, 3, and 4) have lower empirical frequencies. Some consideration sets, including multiple PBMA burgers, have much higher frequencies than the independent logit predictions. This furnishes empirical support for complex consideration. The figure also includes the predictions of a simple multivariate logit model and the version used in the *Main Manuscript* (“full model”) that includes observed heterogeneity and price effects. Both models produce results that closely align with the data (χ^2 -tests are insignificant with χ^2 -values < 15), showing that accounting for pairwise associations between alternatives is crucial for modeling burger consideration.

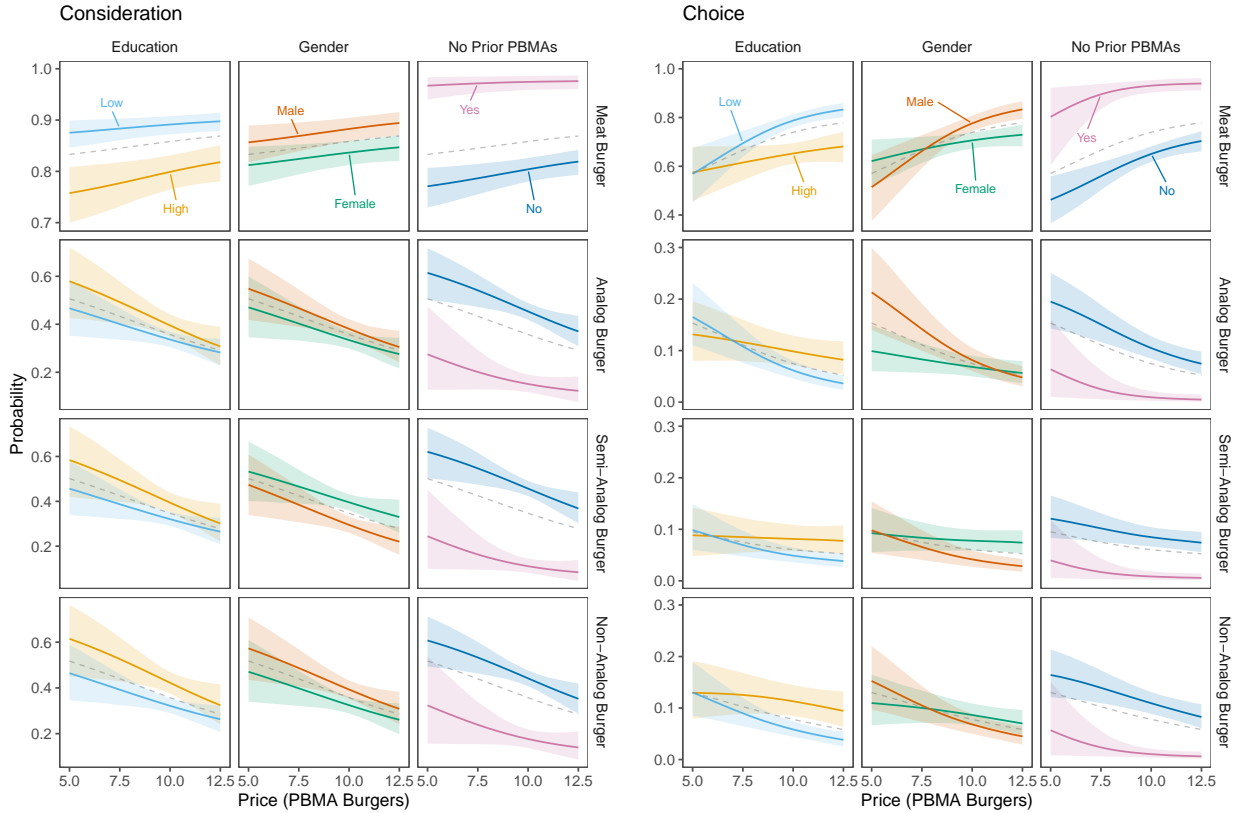


Fig. S5. Price Effect Simulations (Study 2, Part I): Education, Gender, and No Prior PBMA Eaters.

Note: Consideration and choice probabilities are shown for each burger option given a PBMA price range between \$12.50 and \$5, along with 95% credible intervals. Across PBMA price scenarios, the price of the meat burger was kept constant at \$10. The grey dashed lines display the population mean (representative of the United States). No Prior PBMA Eaters (=Yes) are those who self-report to never eat plant-based meat alternatives.

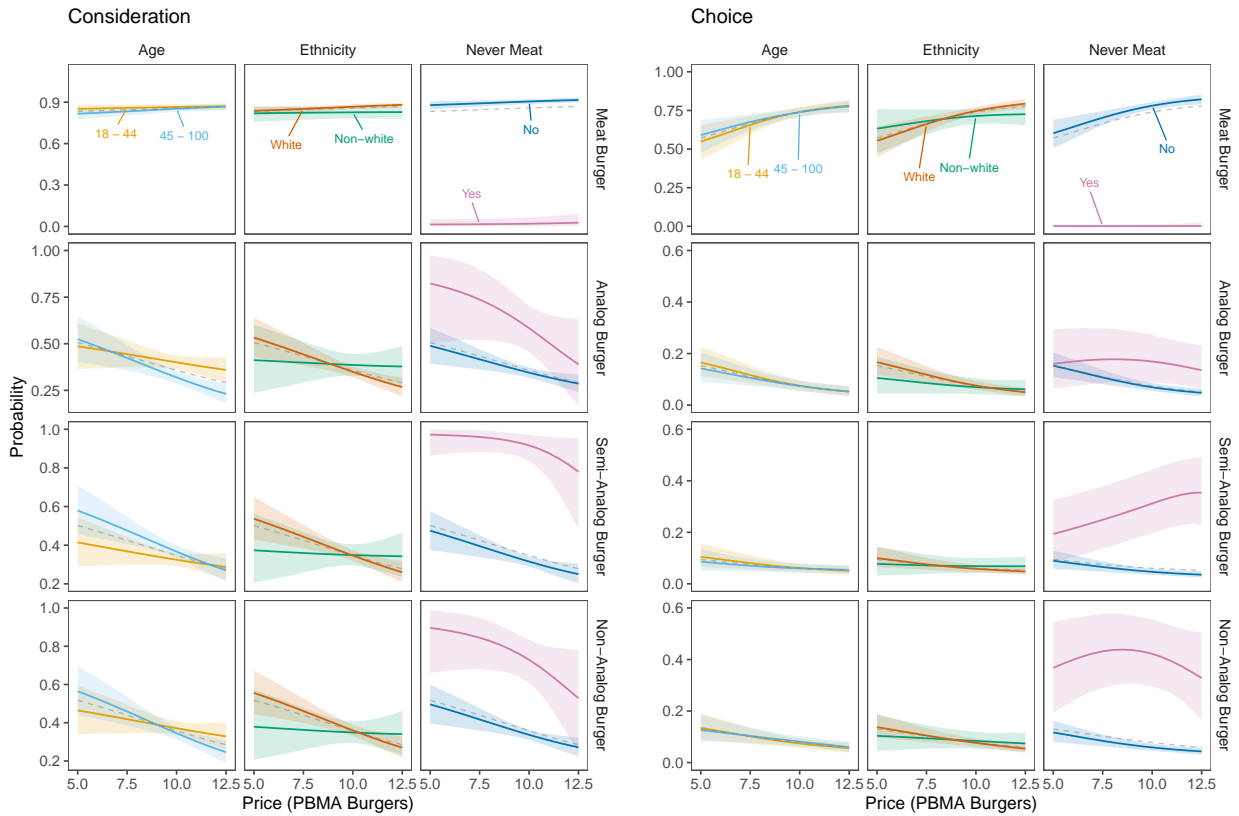


Fig. S6. Price Effect Simulations (Study 2, Part II): Age, Ethnicity, and Never Meat Eaters. Note: Consideration and choice probabilities are shown for each burger option given a PBMA price range between \$12.50 and \$5, along with 95% credible intervals. Across PBMA price scenarios, the price of the meat burger was kept constant at \$10. The grey dashed lines display the population mean (representative of the United States). Never Meat (=Yes) are those who self-report to never eat meat.

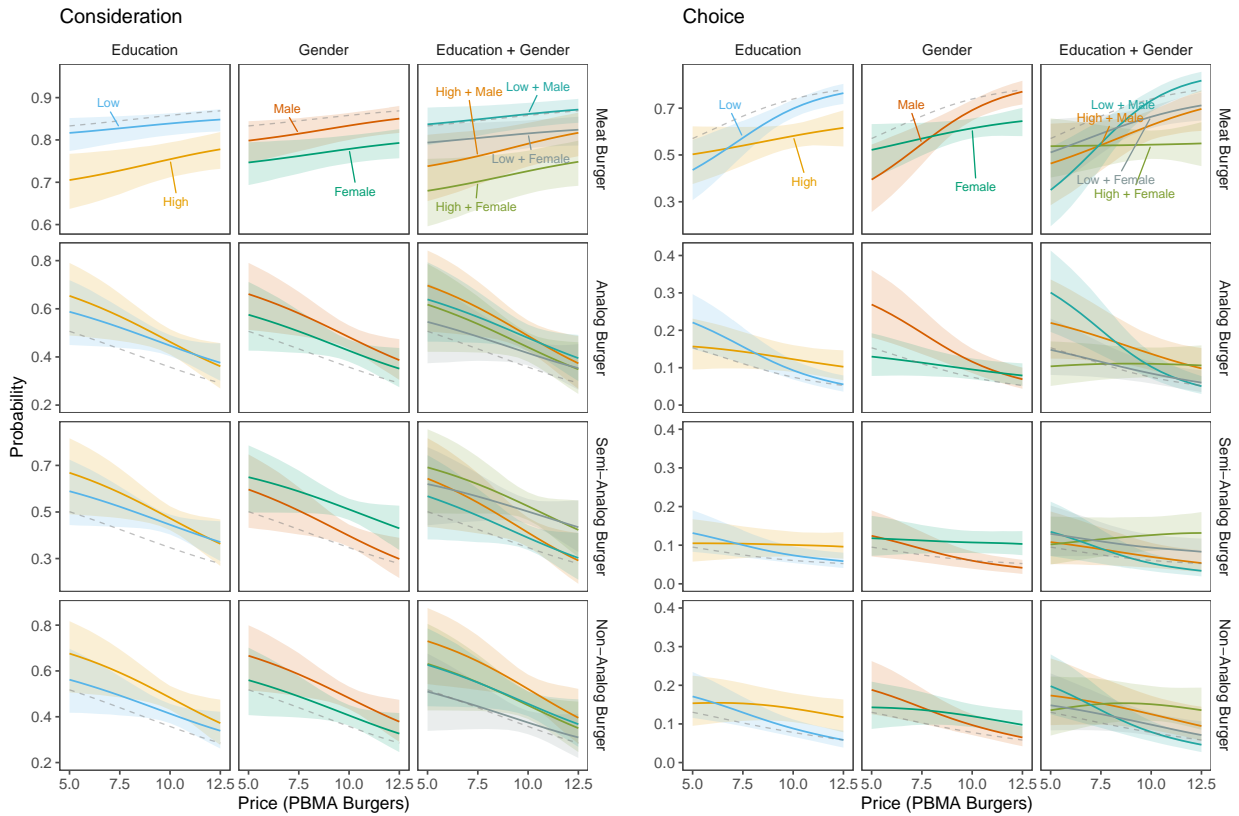


Fig. S7. Price Effect Simulations (Study 2, Part III): Education and Gender Differences After Exclusion of No Prior PBMA Eaters.

Note: Consideration and choice probabilities are shown for each burger option given a PBMA price range between \$12.50 and \$5, along with 95% credible intervals. Across PBMA price scenarios, the price of the meat burger was kept constant at \$10. The grey dashed lines display the population mean (representative of the United States). No Prior PBMA Eaters (=Yes) are those who self-report to never eat plant-based meat alternatives.

Supporting Information: Tables

Table S1. Study 1 Sample Weighting.

Age	Female						Male					
	Less Than Bachelor's Degree			Bachelor's Degree or Higher			Less Than Bachelor's Degree			Bachelor's Degree or Higher		
	Sample %	U.S. %	Weight	Sample %	U.S. %	Weight	Sample %	U.S. %	Weight	Sample %	U.S. %	Weight
18-24	2.7	4.8	1.75	2.9	0.8	0.28	1.7	5.1	2.95	1.2	0.6	0.48
25-44	13.1	9.3	0.69	15.2	7.9	0.51	12.2	10.7	0.86	13.2	7.9	0.48
45-64	9.2	10.1	1.08	7.9	6.3	0.78	4.4	10.4	2.31	5.7	6.3	0.94
Over 65	2.5	8.4	3.31	2.6	3.5	1.33	1.3	6.4	4.84	2.2	3.5	1.62

Note: Participants that could not be assigned a specific cell ($N = 22$) were weighted to fit the sample mean. U.S. % based on 2020 U.S. Census data [27].

Table S2. Average Utility and Observed Heterogeneity (Study 1).

	Intercept	Age	Gender	Education	Ethnicity	Never Meat	No Prior PBMA
Meat Burger	4.13 [3.61, 4.71]	0.02 [-0.14, 0.19]	-0.96 [-1.49, -0.47]	-0.69 [-1.18, -0.21]	0.13 [-0.46, 0.72]	-4.13 [-5.20, -3.13]	1.60 [0.91, 2.36]
Analog Burger	-1.01 [-1.23, -0.80]	-0.10 [-0.21, 0.02]	0.06 [-0.28, 0.40]	0.51 [0.17, 0.87]	0.02 [-0.39, 0.42]	0.99 [0.10, 1.93]	-3.10 [-3.69, -2.57]
Semi-Analog Burger	-0.34 [-0.54, -0.15]	0.07 [-0.04, 0.19]	0.48 [0.14, 0.82]	0.09 [-0.24, 0.44]	-0.04 [-0.47, 0.36]	3.82 [2.52, 5.23]	-2.70 [-3.25, -2.19]
Non-Analog Burger	-0.41 [-0.60, -0.22]	-0.20 [-0.32, -0.08]	0.02 [-0.33, 0.37]	0.47 [0.13, 0.83]	-0.28 [-0.70, 0.15]	0.14 [-0.79, 1.08]	-1.92 [-2.45, -1.43]

Note: Parameter and 95% credible intervals (CI) in parentheses. Significant estimates (at 5%) are bolded.

Table S3. Unobserved Heterogeneity and Scale Factor (Study 1).

Heterogeneity	σ	95% CI
Meat Burger	1.59	[1.23, 1.96]
Analog Burger	1.24	[1.00, 1.52]
Semi-Analog Burger	2.53	[2.14, 2.93]
Non-Analog Burger	1.68	[1.43, 1.96]
Scaling	μ	95% CI
Scale factor	1.11	[0.92, 1.34]

Table S4. Correlations (Study 1).

Alternative	Meat Burger	Analog Burger	Semi-Analog Burger	Non-Analog Burger
Meat Burger	1			
Analog Burger	-0.16 [-0.45, 0.15]	1		
Semi-Analog Burger	-0.39 [-0.51, -0.26]	0.63 [0.51, 0.71]	1	
Non-Analog Burger	-0.24 [-0.43, -0.04]	0.40 [0.21, 0.58]	0.36 [0.27, 0.45]	1

Note: Parameters are shown with 95% credible intervals (CI) in parentheses. Significant estimates (at 5%) are bolded.

Table S5. Market Share Simulations (Study 1).

Scenario	Meat Burger		Analog Burger		Semi-Analog Burger		Non-Analog Burger		Outside Option	
	Mean	95% CI	Mean	95% CI	Mean	95% CI	Mean	95% CI	Mean	95% CI
One Meat Option, No PBMA Options										
1	91.8	[90.8, 92.7]							8.2	[7.3, 9.2]
One Meat Option, One PBMA Option										
2	84.7	[83.6, 85.8]	12.5	[11.4, 13.5]					2.8	[2.1, 3.5]
3	82.2	[81.1, 83.3]			16.1	[15.1, 17.1]			1.7	[1.2, 2.3]
4	82.6	[81.4, 83.9]					14.7	[13.5, 16.0]	2.6	[2.0, 3.3]
One Meat Option, Two PBMA Options										
5	79.6	[78.4, 80.7]	6.2	[5.3, 7.2]	13.2	[12.1, 14.3]			1.0	[0.6, 1.5]
6	79.0	[77.7, 80.2]	8.7	[7.8, 9.7]			11.2	[10.0, 12.4]	1.1	[0.7, 1.6]
7	77.1	[75.8, 78.4]			12.9	[11.9, 13.9]	9.4	[8.2, 10.5]	0.7	[0.4, 1.0]
One Meat Option, Three PBMA Options										
8	75.4	[74.0, 76.7]	4.8	[4.0, 5.6]	11.1	[10.1, 12.2]	8.3	[7.3, 9.4]	0.5	[0.2, 0.7]
No Meat Option, One PBMA Option										
9			37.9	[36.3, 39.5]					62.1	[60.5, 63.7]
10					45.2	[43.7, 46.9]			54.8	[53.1, 56.3]
11							44.9	[43.2, 46.5]	55.1	[53.5, 56.8]
No Meat Option, Two PBMA Options										
12			21.3	[19.6, 23.0]	35.1	[33.2, 37.0]			43.6	[41.9, 45.2]
13			25.0	[23.4, 26.7]			34.0	[32.2, 35.8]	41.0	[39.4, 42.7]
14					33.1	[31.3, 34.8]	30.7	[29.0, 32.4]	36.3	[34.5, 38.0]
No Meat Option, Three PBMA Options										
15			15.0	[13.5, 16.5]	27.3	[25.4, 29.0]	26.6	[24.9, 28.3]	31.2	[29.5, 32.9]

Note: Compared to a scenario with one meat and one PBMA options, adding a second PBMA option increases PBMA share and decreases meat share. For example, the average meat share across scenarios 2–4 is 83.2%, while it is 78.6% across scenarios 5–7, a decrease by 5.6%. Conversely, the average PBMA share increased from 14.4% to 20.5% (+42.4%). When 3 PBMA options are available (scenario 8), meat share further decreases by 4% (to 75.4%), while PBMA share increases by 18.0% (to 24.2%). Comparing scenarios 8 and 15, collective PBMA share increases from 24.2% to 68.9% (+184.7%).

Table S6. Burger Prices at Selected Hamburger Restaurant Chains in the United States.

Hamburger Chain	Meat Burger	Non-Meat Burger
Burger King	\$5.79 (Whopper)	\$6.69 (Impossible Whopper)
Carl's Jr.	\$6.99 (Big Carl's)	Not sold
Fatburger	\$8.99 (Original Fatburger)	\$11.99 (Impossible Burger)
Five Guys	\$10.69 (Hamburger)	Not sold
In-N-Out Burger	\$6.19 (Double Double)	Not sold
Killer Burger	\$10.75 (Classic)	\$10.75 (Classic)
McDonald's	\$4.69 (Big Mac)	\$5.49 (McPlant ^a)
Mooyah	\$8.29 (MOOYAH Cheeseburger)	\$12.29 (The Meatless Beast)
Red Robin	\$13.99 (Keep It Simple)	\$17.49 (Keep It Simple)
Shake Shack	\$7.89 (ShackBurger)	\$9.49 (Veggie Shack)
Smashburger	\$7.79 (Classic Smash Burger)	\$9.59 (Classic Smash Veggie Burger)
Whataburger	\$5.39 (Whataburger)	Not sold

Prices as of July 2024. ^aThe McPlant has been discontinued in the United States.

Table S7. Sample Weighting (Study 2).

Age	Female						Male					
	Less Than Bachelor's Degree			Bachelor's Degree or Higher			Less Than Bachelor's Degree			Bachelor's Degree or Higher		
	Sample %	U.S. %	Weight	Sample %	U.S. %	Weight	Sample %	U.S. %	Weight	Sample %	U.S. %	Weight
18-24	3.5	4.8	1.35	2.0	0.8	0.41	2.1	5.1	2.33	0.7	0.6	0.80
25-44	11.8	9.3	0.76	14.7	7.9	0.51	11.0	10.7	0.94	12.4	7.9	0.51
45-64	10.7	10.1	0.92	9.0	6.3	0.68	5.4	10.4	1.86	6.7	6.3	0.80
Over 65	2.5	8.4	3.29	2.3	3.5	1.48	1.0	6.4	5.00 ^a	1.3	3.5	2.64

Note: Participants that could not be assigned a specific cell ($N = 32$) were weighted to fit the sample mean. U.S. % based on 2020 U.S. Census data [27]. ^aThe estimated weight of 6.367 was capped at 5 [28].

Table S8. Parameter Estimates and Credible Intervals (Study 2).

Consideration	Intercept	Age	Gender	Education	Ethnicity	Never Meat	No Prior PBMA
Meat Burger	3.07 [2.64, 3.52]	-0.14 [-0.28, -0.00]	-0.29 [-0.75, 0.17]	-0.72 [-1.16, -0.28]	0.54 [0.00, 1.05]	-6.39 [-9.16, -4.56]	1.54 [0.75, 2.44]
Analog Burger	-0.71 [-1.13, -0.28]	-0.12 [-0.21, -0.04]	-0.34 [-0.62, -0.08]	0.04 [-0.24, 0.32]	0.02 [-0.31, 0.35]	-0.15 [-0.83, 0.54]	-1.13 [-1.47, -0.79]
Semi-Analog Burger	-1.23 [-1.72, -0.75]	0.18 [0.08, 0.28]	0.69 [0.38, 0.99]	-0.09 [-0.40, 0.22]	-0.05 [-0.42, 0.33]	2.22 [1.20, 3.42]	-1.41 [-1.81, -1.02]
Non-Analog Burger	-0.58 [-1.03, -0.12]	-0.06 [-0.15, 0.03]	-0.60 [-0.90, -0.31]	0.25 [-0.04, 0.54]	0.21 [-0.14, 0.54]	0.13 [-0.60, 0.88]	-0.63 [-0.99, -0.28]
Price (Meat)	0.17 [-0.08, 0.42]	0.02 [-0.08, 0.14]	-0.09 [-0.47, 0.30]	-0.12 [-0.49, 0.23]	-0.58 [-1.01, -0.19]	-0.16 [-1.57, 1.34]	0.09 [-0.59, 0.72]
Price (PBMA)	-0.11 [-0.18, -0.04]	-0.03 [-0.07, 0.01]	0.03 [-0.10, 0.16]	-0.03 [-0.17, 0.11]	-0.09 [-0.26, 0.07]	-0.19 [-0.55, 0.14]	0.01 [-0.16, 0.18]
Choice	Intercept	Age	Gender	Education	Ethnicity	Never Meat	No Prior PBMA
Meat Burger	3.20 [2.38, 3.82]	0.05 [-0.16, 0.27]	-0.86 [-1.59, -0.15]	-0.11 [-0.83, 0.61]	0.26 [-0.52, 1.00]	-8.82 [-23.12, -0.37]	0.07 [-0.67, 0.83]
Analog Burger	1.64 [1.06, 2.20]	0.14 [-0.10, 0.38]	-0.85 [-1.71, -0.03]	0.23 [-0.60, 1.06]	0.63 [-0.29, 1.58]	-0.10 [-1.35, 1.29]	-1.65 [-3.23, -0.30]
Semi-Analog Burger	1.30 [0.71, 1.90]	-0.01 [-0.27, 0.23]	-0.36 [-1.27, 0.50]	0.36 [-0.47, 1.23]	0.30 [-0.63, 1.20]	0.34 [-0.81, 1.65]	-1.221 [-2.67, 0.19]
Non-Analog Burger	1.45 [0.93, 1.97]	0.13 [-0.12, 0.38]	-0.30 [-1.17, 0.56]	0.40 [-0.43, 1.23]	0.23 [-0.69, 1.14]	0.98 [-0.21, 2.32]	-1.64 [-2.95, -0.42]
Price	-0.19 [-0.34, -0.04]	0.08 [0.01, 0.15]	0.29 [0.05, 0.52]	0.35 [0.13, 0.58]	-0.12 [-0.41, 0.15]	0.47 [0.08, 0.88]	-0.13 [-0.52, 0.26]

Note: Significant parameter estimates (at 5%) are bolded. Association parameters are shown in the *Main Manuscript*, Table 1. No Prior PBMA (=Yes) are those who self-report to never eat plant-based meat alternatives. Never Meat (=Yes) are those who self-report to never eat meat.

Supporting Information: References

- [1] Train K (2009) *Discrete Choice Methods with Simulation*. (Cambridge University Press), 2nd edition.
- [2] Chapman R, Staelin R (1982) Exploiting rank ordered choice set data within the stochastic utility model. *Journal of Marketing Research* 19(3):288–301.
- [3] Rossi P, Allenby G, Misra S (2024) *Bayesian Statistics and Marketing*. (Wiley), 2nd edition.
- [4] Dyachenko T, Reczek R, Allenby G (2014) Models of sequential evaluation in best-worst choice tasks. *Marketing Science* 33(6):828–848.
- [5] Chrzan K, Orme B (2019) *Applied MaxDiff: A Practitioner’s Guide to Best-Worst Scaling*. (Sawtooth Software).
- [6] Schramm JB, Lichters M (2024) Incentive alignment in anchored MaxDiff yields superior predictive validity. *Marketing Letters*. Advance online publication: <https://link.springer.com/article/10.1007/s11002-023-09714-2> [Accessed August 4, 2024].
- [7] McElreath R (2020) *Statistical Rethinking: A Bayesian Course with Examples in R and Stan*. (CRC Press), 2nd edition.
- [8] Lewandowski R, Kurowicka D, Joe H (2009) Generating random correlation matrices based on vines and extended onion method. *Journal of Multivariate Analysis* 100(9):1989–2001.
- [9] Akinc D, Vandebroek M (2018) Bayesian estimation of mixed logit models: Selecting an appropriate prior for the covariance matrix. *Journal of Choice Modelling* 29:133–151.
- [10] Carpenter B, et al. (2017) Stan: A probabilistic programming language. *Journal of Statistical Software* 76(1):1–32.
- [11] Hoffman M, Gelman A (2014) The No-U-turn sampler: adaptively setting path lengths in Hamiltonian Monte Carlo. *Journal of Machine Learning Research* 15:1593–1623.
- [12] Stan Development Team (2023) RStan: the R interface to Stan. R package version 2.32.5.
- [13] Gelman A, et al. (2013) *Bayesian Data Analysis*. (CRC Press), 3rd edition.
- [14] Manrai A, Andrews R (1998) Two-stage discrete choice models for scanner panel data: An assessment of process and assumptions. *European Journal of Operational Research* 111(2):193–215.
- [15] Louviere J, Hensher D, Swait J (2000) *Stated Choice Methods*. (Cambridge University Press).

- [16] Swait J (1984) *Probabilistic Choice Set Generation in Transportation Demand Models*. (Ph.D. Thesis, Massachusetts Institute of Technology).
- [17] Abaluck J, Adams-Prassl A (2021) What do consumers consider before they choose? identification from asymmetric demand responses. *Quarterly Journal of Economics* 136(3):1611–1663.
- [18] Chiang J, Chib S, Narasimhan C (1998) Markov chain Monte Carlo and models of consideration set and parameter heterogeneity. *Journal of Econometrics* 89(1–2):223–248.
- [19] Li Y, et al. (2022) The more you ask, the less you get: When additional questions hurt external validity. *Journal of Marketing Research* 59(5):963–982.
- [20] Sablotny-Wackershauser V, Lichters M, Guhl D, Bengart P, Vogt B (2024) Crossing incentive alignment and adaptive designs in choice-based conjoint: A fruitful endeavor. *Journal of the Academy of Marketing Science* 52(3):610–633.
- [21] Aouad A, Farias V, Levi R (2021) Assortment optimization under consider-then-choose choice models. *Management Science* 67(6):3368–3386.
- [22] Manski C (1977) The structure of random utility models. *Theory and Decision* 8:229–254.
- [23] Amano T, Rhodes A, Seiler S (2022) Flexible demand estimation with search data. Available at SSRN: <https://doi.org/10.2139/ssrn.3214812> [Accessed August 4, 2024].
- [24] Besag J (1974) Spatial interaction and the statistical analysis of lattice systems. *Journal of the Royal Statistical Society: Series B (Methodological)* 36(2):192–225.
- [25] Bel K, Fok D, Paap R (2018) Parameter estimation in multivariate logit models with many binary choices. *Econometric Reviews* 37(5):534–550.
- [26] Russell G, Petersen A (2000) Analysis of cross category dependence in market basket selection. *Journal of Retailing* 76(3):367–392.
- [27] United States Census Bureau (2020) Census data. Available at: <https://data.census.gov/> [Accessed November 28, 2023].
- [28] Helveston JP, et al. (2015) Will subsidies drive electric vehicle adoption? Measuring consumer preferences in the u.s. and china. *Transportation Research Part A: Policy and Practice* 73:96–112.